

# 18.175: Lecture 6

## Borel-Cantelli and strong law

Scott Sheffield

MIT

Weak law of large numbers: characteristic function approach

Laws of large numbers: Borel-Cantelli applications

Strong law of large numbers

Weak law of large numbers: characteristic function approach

Laws of large numbers: Borel-Cantelli applications

Strong law of large numbers

# Statement of weak law of large numbers

- ▶ Suppose  $X_i$  are i.i.d. random variables with mean  $\mu$ .

# Statement of weak law of large numbers

- ▶ Suppose  $X_i$  are i.i.d. random variables with mean  $\mu$ .
- ▶ Then the value  $A_n := \frac{X_1 + X_2 + \dots + X_n}{n}$  is called the *empirical average* of the first  $n$  trials.

# Statement of weak law of large numbers

- ▶ Suppose  $X_i$  are i.i.d. random variables with mean  $\mu$ .
- ▶ Then the value  $A_n := \frac{X_1 + X_2 + \dots + X_n}{n}$  is called the *empirical average* of the first  $n$  trials.
- ▶ We'd guess that when  $n$  is large,  $A_n$  is typically close to  $\mu$ .

# Statement of weak law of large numbers

- ▶ Suppose  $X_i$  are i.i.d. random variables with mean  $\mu$ .
- ▶ Then the value  $A_n := \frac{X_1 + X_2 + \dots + X_n}{n}$  is called the *empirical average* of the first  $n$  trials.
- ▶ We'd guess that when  $n$  is large,  $A_n$  is typically close to  $\mu$ .
- ▶ Indeed, **weak law of large numbers** states that for all  $\epsilon > 0$  we have  $\lim_{n \rightarrow \infty} P\{|A_n - \mu| > \epsilon\} = 0$ .

# Statement of weak law of large numbers

- ▶ Suppose  $X_i$  are i.i.d. random variables with mean  $\mu$ .
- ▶ Then the value  $A_n := \frac{X_1 + X_2 + \dots + X_n}{n}$  is called the *empirical average* of the first  $n$  trials.
- ▶ We'd guess that when  $n$  is large,  $A_n$  is typically close to  $\mu$ .
- ▶ Indeed, **weak law of large numbers** states that for all  $\epsilon > 0$  we have  $\lim_{n \rightarrow \infty} P\{|A_n - \mu| > \epsilon\} = 0$ .
- ▶ Example: as  $n$  tends to infinity, the probability of seeing more than  $.50001n$  heads in  $n$  fair coin tosses tends to zero.



## Extent of weak law

- ▶ Question: does the weak law of large numbers apply no matter what the probability distribution for  $X$  is?

## Extent of weak law

- ▶ Question: does the weak law of large numbers apply no matter what the probability distribution for  $X$  is?
- ▶ Is it always the case that if we define  $A_n := \frac{X_1+X_2+\dots+X_n}{n}$  then  $A_n$  is typically close to some fixed value when  $n$  is large?

## Extent of weak law

- ▶ Question: does the weak law of large numbers apply no matter what the probability distribution for  $X$  is?
- ▶ Is it always the case that if we define  $A_n := \frac{X_1+X_2+\dots+X_n}{n}$  then  $A_n$  is typically close to some fixed value when  $n$  is large?
- ▶ What if  $X$  is Cauchy?

## Extent of weak law

- ▶ Question: does the weak law of large numbers apply no matter what the probability distribution for  $X$  is?
- ▶ Is it always the case that if we define  $A_n := \frac{X_1+X_2+\dots+X_n}{n}$  then  $A_n$  is typically close to some fixed value when  $n$  is large?
- ▶ What if  $X$  is Cauchy?
- ▶ In this strange and delightful case  $A_n$  actually has the same probability distribution as  $X$ .

## Extent of weak law

- ▶ Question: does the weak law of large numbers apply no matter what the probability distribution for  $X$  is?
- ▶ Is it always the case that if we define  $A_n := \frac{X_1+X_2+\dots+X_n}{n}$  then  $A_n$  is typically close to some fixed value when  $n$  is large?
- ▶ What if  $X$  is Cauchy?
- ▶ In this strange and delightful case  $A_n$  actually has the same probability distribution as  $X$ .
- ▶ In particular, the  $A_n$  are not tightly concentrated around any particular value even when  $n$  is very large.

## Extent of weak law

- ▶ Question: does the weak law of large numbers apply no matter what the probability distribution for  $X$  is?
- ▶ Is it always the case that if we define  $A_n := \frac{X_1+X_2+\dots+X_n}{n}$  then  $A_n$  is typically close to some fixed value when  $n$  is large?
- ▶ What if  $X$  is Cauchy?
- ▶ In this strange and delightful case  $A_n$  actually has the same probability distribution as  $X$ .
- ▶ In particular, the  $A_n$  are not tightly concentrated around any particular value even when  $n$  is very large.
- ▶ But weak law holds as long as  $E[|X|]$  is finite, so that  $\mu$  is well defined.

## Extent of weak law

- ▶ Question: does the weak law of large numbers apply no matter what the probability distribution for  $X$  is?
- ▶ Is it always the case that if we define  $A_n := \frac{X_1+X_2+\dots+X_n}{n}$  then  $A_n$  is typically close to some fixed value when  $n$  is large?
- ▶ What if  $X$  is Cauchy?
- ▶ In this strange and delightful case  $A_n$  actually has the same probability distribution as  $X$ .
- ▶ In particular, the  $A_n$  are not tightly concentrated around any particular value even when  $n$  is very large.
- ▶ But weak law holds as long as  $E[|X|]$  is finite, so that  $\mu$  is well defined.
- ▶ One standard proof uses characteristic functions.

# Characteristic functions

- ▶ Let  $X$  be a random variable.



# Characteristic functions

- ▶ Let  $X$  be a random variable.
- ▶ The **characteristic function** of  $X$  is defined by  $\phi(t) = \phi_X(t) := E[e^{itX}]$ . Like  $M(t)$  except with  $i$  thrown in.

# Characteristic functions

- ▶ Let  $X$  be a random variable.
- ▶ The **characteristic function** of  $X$  is defined by  $\phi(t) = \phi_X(t) := E[e^{itX}]$ . Like  $M(t)$  except with  $i$  thrown in.
- ▶ Recall that by definition  $e^{it} = \cos(t) + i \sin(t)$ .

# Characteristic functions

- ▶ Let  $X$  be a random variable.
- ▶ The **characteristic function** of  $X$  is defined by  $\phi(t) = \phi_X(t) := E[e^{itX}]$ . Like  $M(t)$  except with  $i$  thrown in.
- ▶ Recall that by definition  $e^{it} = \cos(t) + i \sin(t)$ .
- ▶ Characteristic functions are similar to moment generating functions in some ways.

# Characteristic functions

- ▶ Let  $X$  be a random variable.
- ▶ The **characteristic function** of  $X$  is defined by  $\phi(t) = \phi_X(t) := E[e^{itX}]$ . Like  $M(t)$  except with  $i$  thrown in.
- ▶ Recall that by definition  $e^{it} = \cos(t) + i \sin(t)$ .
- ▶ Characteristic functions are similar to moment generating functions in some ways.
- ▶ For example,  $\phi_{X+Y} = \phi_X \phi_Y$ , just as  $M_{X+Y} = M_X M_Y$ , if  $X$  and  $Y$  are independent.

# Characteristic functions

- ▶ Let  $X$  be a random variable.
- ▶ The **characteristic function** of  $X$  is defined by  $\phi(t) = \phi_X(t) := E[e^{itX}]$ . Like  $M(t)$  except with  $i$  thrown in.
- ▶ Recall that by definition  $e^{it} = \cos(t) + i \sin(t)$ .
- ▶ Characteristic functions are similar to moment generating functions in some ways.
- ▶ For example,  $\phi_{X+Y} = \phi_X \phi_Y$ , just as  $M_{X+Y} = M_X M_Y$ , if  $X$  and  $Y$  are independent.
- ▶ And  $\phi_{aX}(t) = \phi_X(at)$  just as  $M_{aX}(t) = M_X(at)$ .

# Characteristic functions

- ▶ Let  $X$  be a random variable.
- ▶ The **characteristic function** of  $X$  is defined by  $\phi(t) = \phi_X(t) := E[e^{itX}]$ . Like  $M(t)$  except with  $i$  thrown in.
- ▶ Recall that by definition  $e^{it} = \cos(t) + i \sin(t)$ .
- ▶ Characteristic functions are similar to moment generating functions in some ways.
- ▶ For example,  $\phi_{X+Y} = \phi_X \phi_Y$ , just as  $M_{X+Y} = M_X M_Y$ , if  $X$  and  $Y$  are independent.
- ▶ And  $\phi_{aX}(t) = \phi_X(at)$  just as  $M_{aX}(t) = M_X(at)$ .
- ▶ And if  $X$  has an  $m$ th moment then  $E[X^m] = i^m \phi_X^{(m)}(0)$ .

# Characteristic functions

- ▶ Let  $X$  be a random variable.
- ▶ The **characteristic function** of  $X$  is defined by  $\phi(t) = \phi_X(t) := E[e^{itX}]$ . Like  $M(t)$  except with  $i$  thrown in.
- ▶ Recall that by definition  $e^{it} = \cos(t) + i \sin(t)$ .
- ▶ Characteristic functions are similar to moment generating functions in some ways.
- ▶ For example,  $\phi_{X+Y} = \phi_X \phi_Y$ , just as  $M_{X+Y} = M_X M_Y$ , if  $X$  and  $Y$  are independent.
- ▶ And  $\phi_{aX}(t) = \phi_X(at)$  just as  $M_{aX}(t) = M_X(at)$ .
- ▶ And if  $X$  has an  $m$ th moment then  $E[X^m] = i^m \phi_X^{(m)}(0)$ .
- ▶ But characteristic functions have an advantage: they are well defined at all  $t$  for all random variables  $X$ .

# Continuity theorems

- ▶ Let  $X$  be random variable,  $X_n$  a sequence of random variables.



# Continuity theorems

- ▶ Let  $X$  be random variable,  $X_n$  a sequence of random variables.
- ▶ Say  $X_n$  **converge in distribution** or **converge in law** to  $X$  if  $\lim_{n \rightarrow \infty} F_{X_n}(x) = F_X(x)$  at all  $x \in \mathbb{R}$  at which  $F_X$  is continuous.

# Continuity theorems

- ▶ Let  $X$  be random variable,  $X_n$  a sequence of random variables.
- ▶ Say  $X_n$  **converge in distribution** or **converge in law** to  $X$  if  $\lim_{n \rightarrow \infty} F_{X_n}(x) = F_X(x)$  at all  $x \in \mathbb{R}$  at which  $F_X$  is continuous.
- ▶ The weak law of large numbers can be rephrased as the statement that  $A_n$  converges in law to  $\mu$  (i.e., to the random variable that is equal to  $\mu$  with probability one).

# Continuity theorems

- ▶ Let  $X$  be random variable,  $X_n$  a sequence of random variables.
- ▶ Say  $X_n$  **converge in distribution** or **converge in law** to  $X$  if  $\lim_{n \rightarrow \infty} F_{X_n}(x) = F_X(x)$  at all  $x \in \mathbb{R}$  at which  $F_X$  is continuous.
- ▶ The weak law of large numbers can be rephrased as the statement that  $A_n$  converges in law to  $\mu$  (i.e., to the random variable that is equal to  $\mu$  with probability one).
- ▶ **Lévy's continuity theorem (coming later):** if

$$\lim_{n \rightarrow \infty} \phi_{X_n}(t) = \phi_X(t)$$

for all  $t$ , then  $X_n$  converge in law to  $X$ .

# Continuity theorems

- ▶ Let  $X$  be random variable,  $X_n$  a sequence of random variables.
- ▶ Say  $X_n$  **converge in distribution** or **converge in law** to  $X$  if  $\lim_{n \rightarrow \infty} F_{X_n}(x) = F_X(x)$  at all  $x \in \mathbb{R}$  at which  $F_X$  is continuous.
- ▶ The weak law of large numbers can be rephrased as the statement that  $A_n$  converges in law to  $\mu$  (i.e., to the random variable that is equal to  $\mu$  with probability one).
- ▶ **Lévy's continuity theorem (coming later):** if

$$\lim_{n \rightarrow \infty} \phi_{X_n}(t) = \phi_X(t)$$

for all  $t$ , then  $X_n$  converge in law to  $X$ .

- ▶ By this theorem, we can prove weak law of large numbers by showing  $\lim_{n \rightarrow \infty} \phi_{A_n}(t) = \phi_\mu(t) = e^{it\mu}$  for all  $t$ . When  $\mu = 0$ , amounts to showing  $\lim_{n \rightarrow \infty} \phi_{A_n}(t) = 1$  for all  $t$ .
- ▶ **Moment generating analog:** if moment generating functions  $M_{X_n}(t)$  are defined for all  $t$  and  $n$  and, for all  $t$ ,  $\lim_{n \rightarrow \infty} M_{X_n}(t) = M_X(t)$ , then  $X_n$  converge in law to  $X$ .

## Proof sketch for weak law of large numbers, finite mean case

- ▶ As above, let  $X_i$  be i.i.d. instances of random variable  $X$  with mean zero. Write  $A_n := \frac{X_1 + X_2 + \dots + X_n}{n}$ . Weak law of large numbers holds for i.i.d. instances of  $X$  if and only if it holds for i.i.d. instances of  $X - \mu$ . Thus it suffices to prove the weak law in the mean zero case.

# Proof sketch for weak law of large numbers, finite mean case

- ▶ As above, let  $X_i$  be i.i.d. instances of random variable  $X$  with mean zero. Write  $A_n := \frac{X_1 + X_2 + \dots + X_n}{n}$ . Weak law of large numbers holds for i.i.d. instances of  $X$  if and only if it holds for i.i.d. instances of  $X - \mu$ . Thus it suffices to prove the weak law in the mean zero case.
- ▶ Consider the characteristic function  $\phi_X(t) = E[e^{itX}]$ .

# Proof sketch for weak law of large numbers, finite mean case

- ▶ As above, let  $X_i$  be i.i.d. instances of random variable  $X$  with mean zero. Write  $A_n := \frac{X_1 + X_2 + \dots + X_n}{n}$ . Weak law of large numbers holds for i.i.d. instances of  $X$  if and only if it holds for i.i.d. instances of  $X - \mu$ . Thus it suffices to prove the weak law in the mean zero case.
- ▶ Consider the characteristic function  $\phi_X(t) = E[e^{itX}]$ .
- ▶ Since  $E[X] = 0$ , we have  $\phi'_X(0) = E\left[\frac{\partial}{\partial t} e^{itX}\right]_{t=0} = iE[X] = 0$ .

# Proof sketch for weak law of large numbers, finite mean case

- ▶ As above, let  $X_i$  be i.i.d. instances of random variable  $X$  with mean zero. Write  $A_n := \frac{X_1 + X_2 + \dots + X_n}{n}$ . Weak law of large numbers holds for i.i.d. instances of  $X$  if and only if it holds for i.i.d. instances of  $X - \mu$ . Thus it suffices to prove the weak law in the mean zero case.
- ▶ Consider the characteristic function  $\phi_X(t) = E[e^{itX}]$ .
- ▶ Since  $E[X] = 0$ , we have  $\phi'_X(0) = E\left[\frac{\partial}{\partial t} e^{itX}\right]_{t=0} = iE[X] = 0$ .
- ▶ Write  $g(t) = \log \phi_X(t)$  so  $\phi_X(t) = e^{g(t)}$ . Then  $g(0) = 0$  and (by chain rule)  $g'(0) = \lim_{\epsilon \rightarrow 0} \frac{g(\epsilon) - g(0)}{\epsilon} = \lim_{\epsilon \rightarrow 0} \frac{g(\epsilon)}{\epsilon} = 0$ .



# Proof sketch for weak law of large numbers, finite mean case

- ▶ As above, let  $X_i$  be i.i.d. instances of random variable  $X$  with mean zero. Write  $A_n := \frac{X_1 + X_2 + \dots + X_n}{n}$ . Weak law of large numbers holds for i.i.d. instances of  $X$  if and only if it holds for i.i.d. instances of  $X - \mu$ . Thus it suffices to prove the weak law in the mean zero case.
- ▶ Consider the characteristic function  $\phi_X(t) = E[e^{itX}]$ .
- ▶ Since  $E[X] = 0$ , we have  $\phi'_X(0) = E\left[\frac{\partial}{\partial t} e^{itX}\right]_{t=0} = iE[X] = 0$ .
- ▶ Write  $g(t) = \log \phi_X(t)$  so  $\phi_X(t) = e^{g(t)}$ . Then  $g(0) = 0$  and (by chain rule)  $g'(0) = \lim_{\epsilon \rightarrow 0} \frac{g(\epsilon) - g(0)}{\epsilon} = \lim_{\epsilon \rightarrow 0} \frac{g(\epsilon)}{\epsilon} = 0$ .
- ▶ Now  $\phi_{A_n}(t) = \phi_X(t/n)^n = e^{ng(t/n)}$ . Since  $g(0) = g'(0) = 0$  we have  $\lim_{n \rightarrow \infty} ng(t/n) = \lim_{n \rightarrow \infty} t \frac{g(\frac{t}{n})}{\frac{t}{n}} = 0$  if  $t$  is fixed. Thus  $\lim_{n \rightarrow \infty} e^{ng(t/n)} = 1$  for all  $t$ .

# Proof sketch for weak law of large numbers, finite mean case

- ▶ As above, let  $X_i$  be i.i.d. instances of random variable  $X$  with mean zero. Write  $A_n := \frac{X_1 + X_2 + \dots + X_n}{n}$ . Weak law of large numbers holds for i.i.d. instances of  $X$  if and only if it holds for i.i.d. instances of  $X - \mu$ . Thus it suffices to prove the weak law in the mean zero case.
- ▶ Consider the characteristic function  $\phi_X(t) = E[e^{itX}]$ .
- ▶ Since  $E[X] = 0$ , we have  $\phi'_X(0) = E\left[\frac{\partial}{\partial t} e^{itX}\right]_{t=0} = iE[X] = 0$ .
- ▶ Write  $g(t) = \log \phi_X(t)$  so  $\phi_X(t) = e^{g(t)}$ . Then  $g(0) = 0$  and (by chain rule)  $g'(0) = \lim_{\epsilon \rightarrow 0} \frac{g(\epsilon) - g(0)}{\epsilon} = \lim_{\epsilon \rightarrow 0} \frac{g(\epsilon)}{\epsilon} = 0$ .
- ▶ Now  $\phi_{A_n}(t) = \phi_X(t/n)^n = e^{ng(t/n)}$ . Since  $g(0) = g'(0) = 0$  we have  $\lim_{n \rightarrow \infty} ng(t/n) = \lim_{n \rightarrow \infty} t \frac{g(\frac{t}{n})}{\frac{t}{n}} = 0$  if  $t$  is fixed. Thus  $\lim_{n \rightarrow \infty} e^{ng(t/n)} = 1$  for all  $t$ .

# Proof sketch for weak law of large numbers, finite mean case

- ▶ As above, let  $X_i$  be i.i.d. instances of random variable  $X$  with mean zero. Write  $A_n := \frac{X_1 + X_2 + \dots + X_n}{n}$ . Weak law of large numbers holds for i.i.d. instances of  $X$  if and only if it holds for i.i.d. instances of  $X - \mu$ . Thus it suffices to prove the weak law in the mean zero case.
- ▶ Consider the characteristic function  $\phi_X(t) = E[e^{itX}]$ .
- ▶ Since  $E[X] = 0$ , we have  $\phi'_X(0) = E\left[\frac{\partial}{\partial t} e^{itX}\right]_{t=0} = iE[X] = 0$ .
- ▶ Write  $g(t) = \log \phi_X(t)$  so  $\phi_X(t) = e^{g(t)}$ . Then  $g(0) = 0$  and (by chain rule)  $g'(0) = \lim_{\epsilon \rightarrow 0} \frac{g(\epsilon) - g(0)}{\epsilon} = \lim_{\epsilon \rightarrow 0} \frac{g(\epsilon)}{\epsilon} = 0$ .
- ▶ Now  $\phi_{A_n}(t) = \phi_X(t/n)^n = e^{ng(t/n)}$ . Since  $g(0) = g'(0) = 0$  we have  $\lim_{n \rightarrow \infty} ng(t/n) = \lim_{n \rightarrow \infty} t \frac{g(t/n)}{t/n} = 0$  if  $t$  is fixed. Thus  $\lim_{n \rightarrow \infty} e^{ng(t/n)} = 1$  for all  $t$ .
- ▶ By Lévy's continuity theorem, the  $A_n$  converge in law to 0 (i.e., to the random variable that is 0 with probability one).

Weak law of large numbers: characteristic function approach

Laws of large numbers: Borel-Cantelli applications

Strong law of large numbers

Weak law of large numbers: characteristic function approach

Laws of large numbers: Borel-Cantelli applications

Strong law of large numbers

- ▶ **First Borel-Cantelli lemma:** If  $\sum_{n=1}^{\infty} P(A_n) < \infty$  then  $P(A_n \text{ i.o.}) = 0$ .

- ▶ **First Borel-Cantelli lemma:** If  $\sum_{n=1}^{\infty} P(A_n) < \infty$  then  $P(A_n \text{ i.o.}) = 0$ .
- ▶ **Second Borel-Cantelli lemma:** If  $A_n$  are independent, then  $\sum_{n=1}^{\infty} P(A_n) = \infty$  implies  $P(A_n \text{ i.o.}) = 1$ .

- ▶ **Theorem:**  $X_n \rightarrow X$  in probability if and only if for every subsequence of the  $X_n$  there is a further subsequence converging a.s. to  $X$ .



# Convergence in probability $\Rightarrow$ subsequential a.s. convergence

- ▶ **Theorem:**  $X_n \rightarrow X$  in probability if and only if for every subsequence of the  $X_n$  there is a further subsequence converging a.s. to  $X$ .
- ▶ **Main idea of proof:** Consider event  $E_n$  that  $X_n$  and  $X$  differ by  $\epsilon$ . Do the  $E_n$  occur i.o.? Use Borel-Cantelli.

## Pairwise independence example

- ▶ **Theorem:** Suppose  $A_1, A_2, \dots$  are pairwise independent and  $\sum P(A_n) = \infty$ , and write  $S_n = \sum_{i=1}^n 1_{A_i}$ . Then the ratio  $S_n/ES_n$  tends a.s. to 1.

## Pairwise independence example

- ▶ **Theorem:** Suppose  $A_1, A_2, \dots$  are pairwise independent and  $\sum P(A_n) = \infty$ , and write  $S_n = \sum_{i=1}^n 1_{A_i}$ . Then the ratio  $S_n/ES_n$  tends a.s. to 1.
- ▶ **Main idea of proof:** First, pairwise independence implies that variances add. Conclude (by checking term by term) that  $\text{Var}S_n \leq ES_n$ . Then Chebyshev implies

$$P(|S_n - ES_n| > \delta ES_n) \leq \text{Var}(S_n)/(\delta ES_n)^2 \rightarrow 0,$$

which gives us convergence in probability.

## Pairwise independence example

- ▶ **Theorem:** Suppose  $A_1, A_2, \dots$  are pairwise independent and  $\sum P(A_n) = \infty$ , and write  $S_n = \sum_{i=1}^n 1_{A_i}$ . Then the ratio  $S_n/ES_n$  tends a.s. to 1.
- ▶ **Main idea of proof:** First, pairwise independence implies that variances add. Conclude (by checking term by term) that  $\text{Var}S_n \leq ES_n$ . Then Chebyshev implies

$$P(|S_n - ES_n| > \delta ES_n) \leq \text{Var}(S_n)/(\delta ES_n)^2 \rightarrow 0,$$

which gives us convergence in probability.

- ▶ Second, take a smart subsequence. Let  $n_k = \inf\{n : ES_n \geq k^2\}$ . Use Borel Cantelli to get a.s. convergence along this subsequence. Check that convergence along this subsequence deterministically implies the non-subsequential convergence.

Weak law of large numbers: characteristic function approach

Laws of large numbers: Borel-Cantelli applications

Strong law of large numbers

Weak law of large numbers: characteristic function approach

Laws of large numbers: Borel-Cantelli applications

Strong law of large numbers

- ▶ **Theorem (strong law):** If  $X_1, X_2, \dots$  are i.i.d. real-valued random variables with expectation  $m$  and  $A_n := n^{-1} \sum_{i=1}^n X_i$  are the *empirical means* then  $\lim_{n \rightarrow \infty} A_n = m$  almost surely.

## Proof of strong law assuming $E[X^4] < \infty$

- ▶ Assume  $K := E[X^4] < \infty$ . Not necessary, but simplifies proof.



## Proof of strong law assuming $E[X^4] < \infty$

- ▶ Assume  $K := E[X^4] < \infty$ . Not necessary, but simplifies proof.
- ▶ Note:  $\text{Var}[X^2] = E[X^4] - E[X^2]^2 \geq 0$ , so  $E[X^2]^2 \leq K$ .

## Proof of strong law assuming $E[X^4] < \infty$

- ▶ Assume  $K := E[X^4] < \infty$ . Not necessary, but simplifies proof.
- ▶ Note:  $\text{Var}[X^2] = E[X^4] - E[X^2]^2 \geq 0$ , so  $E[X^2]^2 \leq K$ .
- ▶ The strong law holds for i.i.d. copies of  $X$  if and only if it holds for i.i.d. copies of  $X - \mu$  where  $\mu$  is a constant.

## Proof of strong law assuming $E[X^4] < \infty$

- ▶ Assume  $K := E[X^4] < \infty$ . Not necessary, but simplifies proof.
- ▶ Note:  $\text{Var}[X^2] = E[X^4] - E[X^2]^2 \geq 0$ , so  $E[X^2]^2 \leq K$ .
- ▶ The strong law holds for i.i.d. copies of  $X$  if and only if it holds for i.i.d. copies of  $X - \mu$  where  $\mu$  is a constant.
- ▶ So we may as well assume  $E[X] = 0$ .

## Proof of strong law assuming $E[X^4] < \infty$

- ▶ Assume  $K := E[X^4] < \infty$ . Not necessary, but simplifies proof.
- ▶ Note:  $\text{Var}[X^2] = E[X^4] - E[X^2]^2 \geq 0$ , so  $E[X^2]^2 \leq K$ .
- ▶ The strong law holds for i.i.d. copies of  $X$  if and only if it holds for i.i.d. copies of  $X - \mu$  where  $\mu$  is a constant.
- ▶ So we may as well assume  $E[X] = 0$ .
- ▶ Key to proof is to bound fourth moments of  $A_n$ .

## Proof of strong law assuming $E[X^4] < \infty$

- ▶ Assume  $K := E[X^4] < \infty$ . Not necessary, but simplifies proof.
- ▶ Note:  $\text{Var}[X^2] = E[X^4] - E[X^2]^2 \geq 0$ , so  $E[X^2]^2 \leq K$ .
- ▶ The strong law holds for i.i.d. copies of  $X$  if and only if it holds for i.i.d. copies of  $X - \mu$  where  $\mu$  is a constant.
- ▶ So we may as well assume  $E[X] = 0$ .
- ▶ Key to proof is to bound fourth moments of  $A_n$ .
- ▶  $E[A_n^4] = n^{-4} E[S_n^4] = n^{-4} E[(X_1 + X_2 + \dots + X_n)^4]$ .

## Proof of strong law assuming $E[X^4] < \infty$

- ▶ Assume  $K := E[X^4] < \infty$ . Not necessary, but simplifies proof.
- ▶ Note:  $\text{Var}[X^2] = E[X^4] - E[X^2]^2 \geq 0$ , so  $E[X^2]^2 \leq K$ .
- ▶ The strong law holds for i.i.d. copies of  $X$  if and only if it holds for i.i.d. copies of  $X - \mu$  where  $\mu$  is a constant.
- ▶ So we may as well assume  $E[X] = 0$ .
- ▶ Key to proof is to bound fourth moments of  $A_n$ .
- ▶  $E[A_n^4] = n^{-4} E[S_n^4] = n^{-4} E[(X_1 + X_2 + \dots + X_n)^4]$ .
- ▶ Expand  $(X_1 + \dots + X_n)^4$ . Five kinds of terms:  $X_i X_j X_k X_l$  and  $X_i X_j X_k^2$  and  $X_i X_j^3$  and  $X_i^2 X_j^2$  and  $X_i^4$ .

## Proof of strong law assuming $E[X^4] < \infty$

- ▶ Assume  $K := E[X^4] < \infty$ . Not necessary, but simplifies proof.
- ▶ Note:  $\text{Var}[X^2] = E[X^4] - E[X^2]^2 \geq 0$ , so  $E[X^2]^2 \leq K$ .
- ▶ The strong law holds for i.i.d. copies of  $X$  if and only if it holds for i.i.d. copies of  $X - \mu$  where  $\mu$  is a constant.
- ▶ So we may as well assume  $E[X] = 0$ .
- ▶ Key to proof is to bound fourth moments of  $A_n$ .
- ▶  $E[A_n^4] = n^{-4} E[S_n^4] = n^{-4} E[(X_1 + X_2 + \dots + X_n)^4]$ .
- ▶ Expand  $(X_1 + \dots + X_n)^4$ . Five kinds of terms:  $X_i X_j X_k X_l$  and  $X_i X_j X_k^2$  and  $X_i X_j^3$  and  $X_i^2 X_j^2$  and  $X_i^4$ .
- ▶ The first three terms all have expectation zero. There are  $\binom{n}{2}$  of the fourth type and  $n$  of the last type, each equal to at most  $K$ . So  $E[A_n^4] \leq n^{-4} \left( 6 \binom{n}{2} + n \right) K$ .

## Proof of strong law assuming $E[X^4] < \infty$

- ▶ Assume  $K := E[X^4] < \infty$ . Not necessary, but simplifies proof.
- ▶ Note:  $\text{Var}[X^2] = E[X^4] - E[X^2]^2 \geq 0$ , so  $E[X^2]^2 \leq K$ .
- ▶ The strong law holds for i.i.d. copies of  $X$  if and only if it holds for i.i.d. copies of  $X - \mu$  where  $\mu$  is a constant.
- ▶ So we may as well assume  $E[X] = 0$ .
- ▶ Key to proof is to bound fourth moments of  $A_n$ .
- ▶  $E[A_n^4] = n^{-4} E[S_n^4] = n^{-4} E[(X_1 + X_2 + \dots + X_n)^4]$ .
- ▶ Expand  $(X_1 + \dots + X_n)^4$ . Five kinds of terms:  $X_i X_j X_k X_l$  and  $X_i X_j X_k^2$  and  $X_i X_j^3$  and  $X_i^2 X_j^2$  and  $X_i^4$ .
- ▶ The first three terms all have expectation zero. There are  $\binom{n}{2}$  of the fourth type and  $n$  of the last type, each equal to at most  $K$ . So  $E[A_n^4] \leq n^{-4} \left( 6 \binom{n}{2} + n \right) K$ .
- ▶ Thus  $E[\sum_{n=1}^{\infty} A_n^4] = \sum_{n=1}^{\infty} E[A_n^4] < \infty$ . So  $\sum_{n=1}^{\infty} A_n^4 < \infty$  (and hence  $A_n \rightarrow 0$ ) with probability 1.



## General proof of strong law

- ▶ Suppose  $X_k$  are i.i.d. with finite mean. Let  $Y_k = X_k 1_{|X_k| \leq k}$ . Write  $T_n = Y_1 + \dots + Y_n$ . **Claim:**  $X_k = Y_k$  all but finitely often a.s. so suffices to show  $T_n/n \rightarrow \mu$ . (Borel Cantelli, expectation of positive r.v. is area between cdf and line  $y = 1$ )

## General proof of strong law

- ▶ Suppose  $X_k$  are i.i.d. with finite mean. Let  $Y_k = X_k 1_{|X_k| \leq k}$ . Write  $T_n = Y_1 + \dots + Y_n$ . **Claim:**  $X_k = Y_k$  all but finitely often a.s. so suffices to show  $T_n/n \rightarrow \mu$ . (Borel Cantelli, expectation of positive r.v. is area between cdf and line  $y = 1$ )
- ▶ **Claim:**  $\sum_{k=1}^{\infty} \text{Var}(Y_k)/k^2 \leq 4E|X_1| < \infty$ . How to prove it?

## General proof of strong law

- ▶ Suppose  $X_k$  are i.i.d. with finite mean. Let  $Y_k = X_k 1_{|X_k| \leq k}$ . Write  $T_n = Y_1 + \dots + Y_n$ . **Claim:**  $X_k = Y_k$  all but finitely often a.s. so suffices to show  $T_n/n \rightarrow \mu$ . (Borel Cantelli, expectation of positive r.v. is area between cdf and line  $y = 1$ )
- ▶ **Claim:**  $\sum_{k=1}^{\infty} \text{Var}(Y_k)/k^2 \leq 4E|X_1| < \infty$ . How to prove it?
- ▶ **Observe:**  $\text{Var}(Y_k) \leq E(Y_k^2) = \int_0^{\infty} 2yP(|Y_k| > y)dy \leq \int_0^k 2yP(|X_1| > y)dy$ . Use Fubini (interchange sum/integral, since everything positive)

$$\sum_{k=1}^{\infty} E(Y_k^2)/k^2 \leq \sum_{k=1}^{\infty} k^{-2} \int_0^{\infty} 1_{(y < k)} 2yP(|X_1| > y)dy = \int_0^{\infty} \left( \sum_{k=1}^{\infty} k^{-2} 1_{(y < k)} \right) 2yP(|X_1| > y)dy.$$

Since  $E|X_1| = \int_0^{\infty} P(|X_1| > y)dy$ , complete proof of claim by showing that if  $y \geq 0$  then  $2y \sum_{k > y} k^{-2} \leq 4$ .

## General proof of strong law

- ▶ **Claim:**  $\sum_{k=1}^{\infty} \text{Var}(Y_k)/k^2 \leq 4E|X_1| < \infty$ . How to use it?

## General proof of strong law

- ▶ **Claim:**  $\sum_{k=1}^{\infty} \text{Var}(Y_k)/k^2 \leq 4E|X_1| < \infty$ . How to use it?
- ▶ Consider subsequence  $k(n) = [\alpha^n]$  for arbitrary  $\alpha > 1$ . Using Chebyshev, if  $\epsilon > 0$  then

$$\begin{aligned} \sum_{n=1}^{\infty} P(|T_{k(n)} - ET_{k(n)}| > \epsilon k(n)) &\leq \epsilon^{-1} \sum_{n=1}^{\infty} \text{Var}(T_{k(n)})/k(n)^2 \\ &= \epsilon^{-2} \sum_{n=1}^{\infty} k(n)^{-2} \sum_{m=1}^{k(n)} \text{Var}(Y_m) = \epsilon^{-2} \sum_{m=1}^{\infty} \text{Var}(Y_m) \sum_{n:k(n) \geq m} k(n)^{-2}. \end{aligned}$$

# General proof of strong law

- ▶ **Claim:**  $\sum_{k=1}^{\infty} \text{Var}(Y_k)/k^2 \leq 4E|X_1| < \infty$ . How to use it?
- ▶ Consider subsequence  $k(n) = [\alpha^n]$  for arbitrary  $\alpha > 1$ . Using Chebyshev, if  $\epsilon > 0$  then

$$\begin{aligned} \sum_{n=1}^{\infty} P(|T_{k(n)} - ET_{k(n)}| > \epsilon k(n)) &\leq \epsilon^{-1} \sum_{n=1}^{\infty} \text{Var}(T_{k(n)})/k(n)^2 \\ &= \epsilon^{-2} \sum_{n=1}^{\infty} k(n)^{-2} \sum_{m=1}^{k(n)} \text{Var}(Y_m) = \epsilon^{-2} \sum_{m=1}^{\infty} \text{Var}(Y_m) \sum_{n:k(n) \geq m} k(n)^{-2}. \end{aligned}$$

- ▶ **Sum series:**

$$\sum_{n:\alpha^n \geq m} [\alpha^n]^{-2} \leq 4 \sum_{n:\alpha^n \geq m} \alpha^{-2n} \leq 4(1 - \alpha^{-2})^{-1} m^{-2}.$$

# General proof of strong law

- ▶ **Claim:**  $\sum_{k=1}^{\infty} \text{Var}(Y_k)/k^2 \leq 4E|X_1| < \infty$ . How to use it?
- ▶ Consider subsequence  $k(n) = [\alpha^n]$  for arbitrary  $\alpha > 1$ . Using Chebyshev, if  $\epsilon > 0$  then

$$\begin{aligned} \sum_{n=1}^{\infty} P(|T_{k(n)} - ET_{k(n)}| > \epsilon k(n)) &\leq \epsilon^{-1} \sum_{n=1}^{\infty} \text{Var}(T_{k(n)})/k(n)^2 \\ &= \epsilon^{-2} \sum_{n=1}^{\infty} k(n)^{-2} \sum_{m=1}^{k(n)} \text{Var}(Y_m) = \epsilon^{-2} \sum_{m=1}^{\infty} \text{Var}(Y_m) \sum_{n:k(n) \geq m} k(n)^{-2}. \end{aligned}$$

- ▶ **Sum series:**

$$\sum_{n:\alpha^n \geq m} [\alpha^n]^{-2} \leq 4 \sum_{n:\alpha^n \geq m} \alpha^{-2n} \leq 4(1 - \alpha^{-2})^{-1} m^{-2}.$$

- ▶ **Combine computations (observe RHS below is finite):**

$$\sum_{n=1}^{\infty} P(|T_{k(n)} - ET_{k(n)}| > \epsilon k(n)) \leq 4(1 - \alpha^{-2})^{-1} \epsilon^{-2} \sum_{m=1}^{\infty} E(Y_m^2) m^{-2}.$$

# General proof of strong law

- ▶ **Claim:**  $\sum_{k=1}^{\infty} \text{Var}(Y_k)/k^2 \leq 4E|X_1| < \infty$ . How to use it?
- ▶ Consider subsequence  $k(n) = [\alpha^n]$  for arbitrary  $\alpha > 1$ . Using Chebyshev, if  $\epsilon > 0$  then

$$\begin{aligned} \sum_{n=1}^{\infty} P(|T_{k(n)} - ET_{k(n)}| > \epsilon k(n)) &\leq \epsilon^{-1} \sum_{n=1}^{\infty} \text{Var}(T_{k(n)})/k(n)^2 \\ &= \epsilon^{-2} \sum_{n=1}^{\infty} k(n)^{-2} \sum_{m=1}^{k(n)} \text{Var}(Y_m) = \epsilon^{-2} \sum_{m=1}^{\infty} \text{Var}(Y_m) \sum_{n:k(n) \geq m} k(n)^{-2}. \end{aligned}$$

- ▶ **Sum series:**

$$\sum_{n:\alpha^n \geq m} [\alpha^n]^{-2} \leq 4 \sum_{n:\alpha^n \geq m} \alpha^{-2n} \leq 4(1 - \alpha^{-2})^{-1} m^{-2}.$$

- ▶ **Combine computations (observe RHS below is finite):**

$$\sum_{n=1}^{\infty} P(|T_{k(n)} - ET_{k(n)}| > \epsilon k(n)) \leq 4(1 - \alpha^{-2})^{-1} \epsilon^{-2} \sum_{m=1}^{\infty} E(Y_m^2) m^{-2}.$$

- ▶ Since  $\epsilon$  is arbitrary, get  $(T_{k(n)} - ET_{k(n)})/k(n) \rightarrow 0$  a.s.



## General proof of strong law

- ▶ Conclude by taking  $\alpha \rightarrow 1$ . This finishes the case that the  $X_1$  are a.s. positive.

# General proof of strong law

- ▶ Conclude by taking  $\alpha \rightarrow 1$ . This finishes the case that the  $X_1$  are a.s. positive.
- ▶ Can extend to the case that  $X_1$  is a.s. positive within infinite mean.

## General proof of strong law

- ▶ Conclude by taking  $\alpha \rightarrow 1$ . This finishes the case that the  $X_1$  are a.s. positive.
- ▶ Can extend to the case that  $X_1$  is a.s. positive within infinite mean.
- ▶ Generally, can consider  $X_1^+$  and  $X_1^-$ , and it is enough if one of them has a finite mean.