



# **Increasing and Decreasing Subsequences**

**Richard P. Stanley**

**M.I.T.**

# Definitions

**3** 1 8 **4** 9 **6** **7** 2 5 (i.s)

3 1 **8** **4** 9 6 7 **2** 5 (d.s)

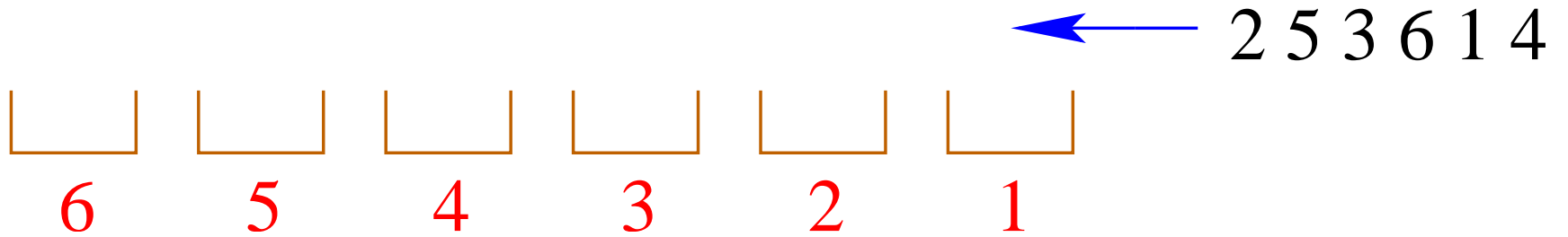
$$\mathbf{is}(w) = |\text{longest i.s.}| = 4$$

$$\mathbf{ds}(w) = |\text{longest d.s.}| = 3$$

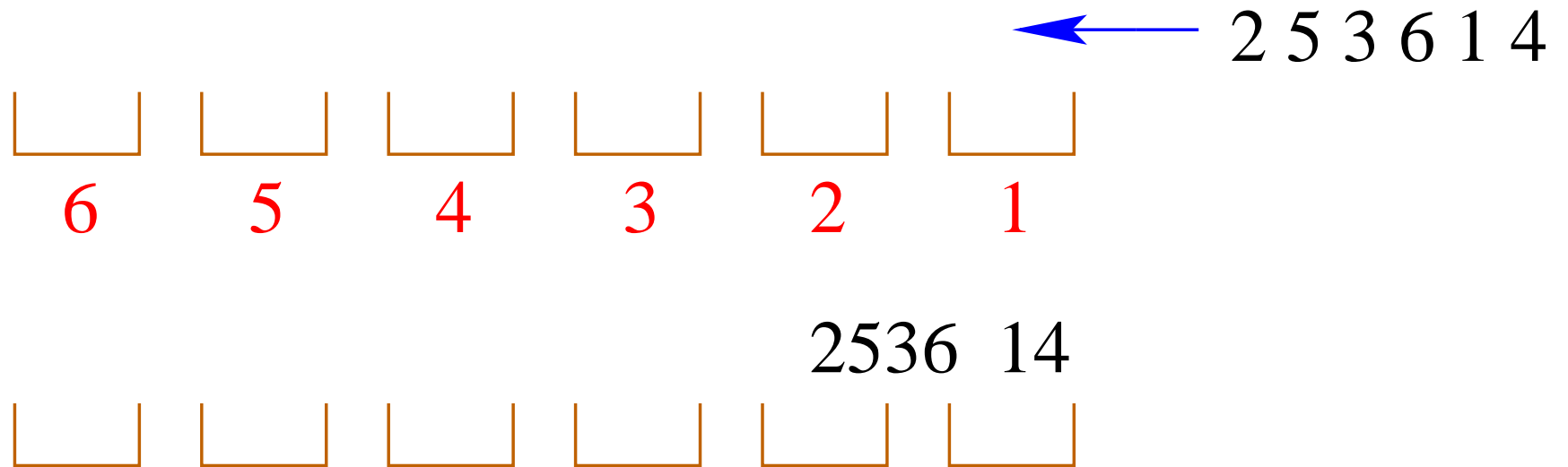
# Application: airplane boarding

**Naive model:** passengers board in order  $w = a_1 a_2 \cdots a_n$  for seats  $1, 2, \dots, n$ . Each passenger takes one time unit to be seated after arriving at his seat.

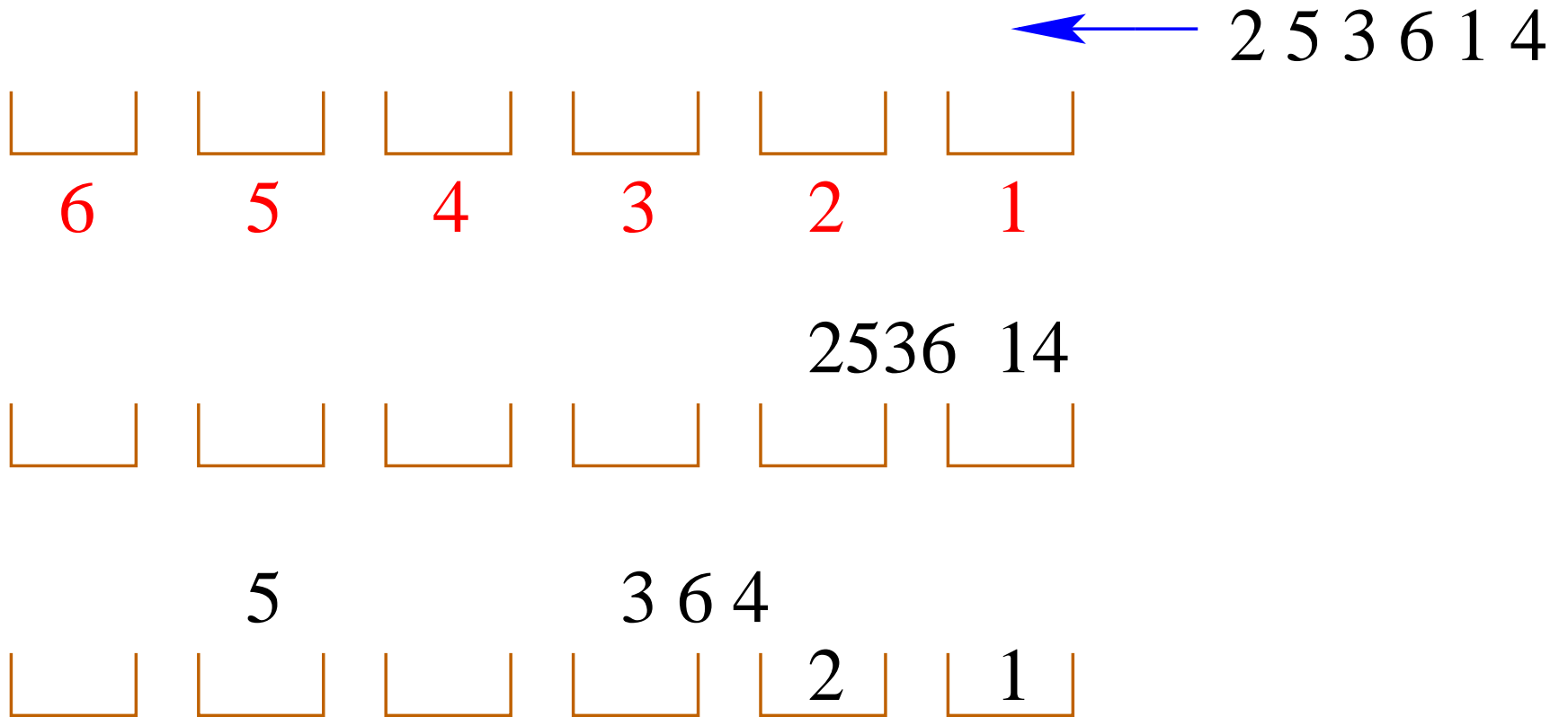
# Boarding process



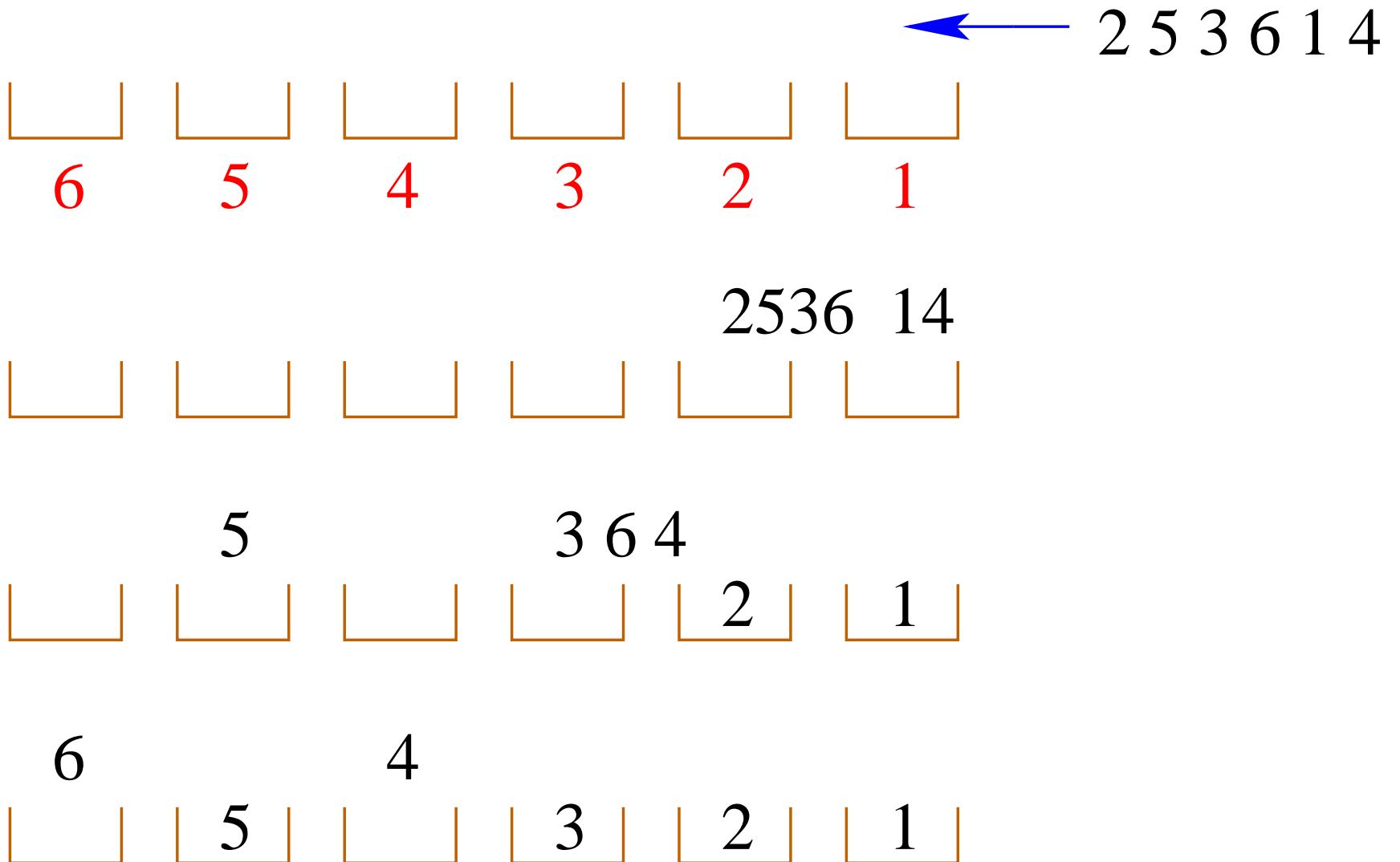
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# Results

**Easy:** Total waiting time =  $is(w)$ .

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**Two conclusions:**

- Usual system (back-to-front) not much better than random.
- Better: first board window seats, then center, then aisle.

# Partitions

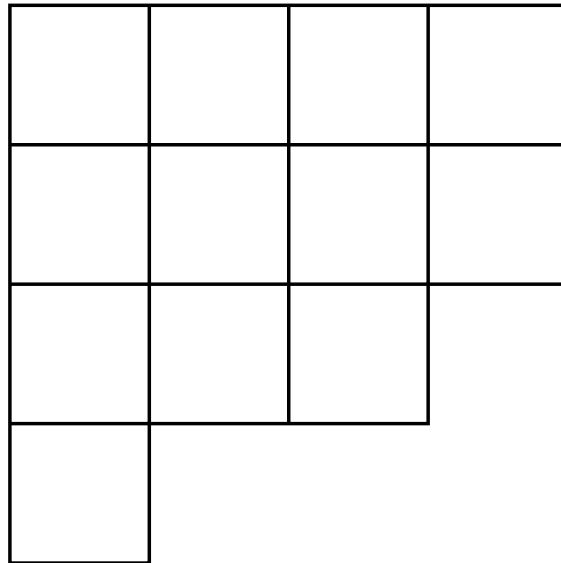
**partition**  $\lambda \vdash n$ :  $\lambda = (\lambda_1, \lambda_2, \dots)$

$$\lambda_1 \geq \lambda_2 \geq \dots \geq 0$$

$$\sum \lambda_i = n$$

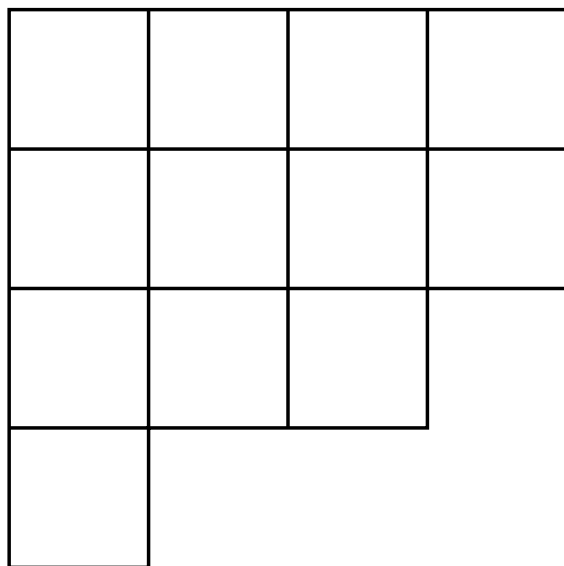
# Young diagrams

(Young) diagram of  $\lambda = (4, 4, 3, 1)$ :

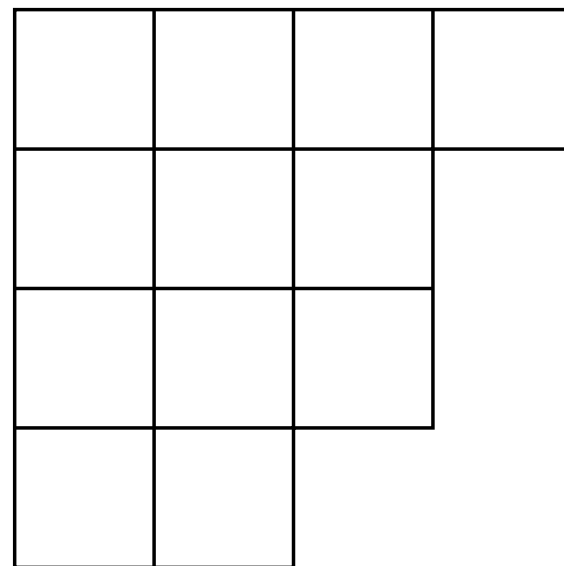


# Conjugate partitions

$\lambda' = (4, 3, 3, 2)$ , the **conjugate** partition to  
 $\lambda = (4, 4, 3, 2)$



$\lambda$



$\lambda'$

# Standard Young tableau

**standard Young tableau** (SYT) of shape  $\lambda \vdash n$ ,  
e.g.,  $\lambda = (4, 4, 3, 1)$ :

<

1	2	7	10
3	5	8	12
4	6	11	
9			

^

$f^\lambda$

$f^\lambda = \#$  of SYT of shape  $\lambda$

E.g.,  $f^{(3,2)} = 5$ :

1 2 3	1 2 4	1 2 5	1 3 4	1 3 5
4 5	3 5	3 4	2 5	2 4

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1 2 3	1 2 4	1 2 5	1 3 4	1 3 5
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$\exists$  simple formula for  $f^\lambda$  (Frame-Robinson-Thrall  
**hook-length formula**)

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**Note.**  $f^\lambda = \dim(\text{irrep. of } \mathfrak{S}_n)$ , where  $\mathfrak{S}_n$  is the  
**symmetric group** of all permutations of  
 $1, 2, \dots, n$ .



# RSK algorithm

**RSK algorithm:** a bijection

$$w \xrightarrow{\text{rsk}} (P, Q),$$

where  $w \in \mathfrak{S}_n$  and  $P, Q$  are SYT of the same shape  $\lambda \vdash n$ .

Write  $\lambda = \mathbf{sh}(w)$ , the **shape** of  $w$ .

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[ea.ea.home.mindspring.com](http://ea.ea.home.mindspring.com)

# Example of RSK: $w = 4132$

insert 4, record 1:	4	1
insert 1, record 2:	1 4	1 2
insert 3, record 3:	1 3 4	1 3 2
insert 2, record 4:	1 2 3 4	1 3 2 4

# Example of RSK: $w = 4132$

insert 4, record 1: 4      1

insert 1, record 2: 1      1  
                          4      2

insert 3, record 3: 1 3    1 3  
                          4      2

insert 2, record 4: 1 2    1 3  
                          3      2  
                          4      4

$$(P, Q) = \left( \begin{array}{c} 1\ 2 \\ 3 \\ 4 \end{array}, \begin{array}{c} 1\ 3 \\ 2 \\ 4 \end{array} \right)$$

# Schensted's theorem

**Theorem.** Let  $w \xrightarrow{\text{rsk}} (P, Q)$ , where  $\text{sh}(P) = \text{sh}(Q) = \lambda$ . Then

$$\text{is}(w) = \text{longest row length} = \lambda_1$$

$$\text{ds}(w) = \text{longest column length} = \lambda'_1.$$

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**Example.**  $4132 \xrightarrow{\text{rsk}} \left( \begin{array}{cc} 1 & 2 \\ 3 & \\ 4 & \end{array} , \begin{array}{cc} 1 & 3 \\ 2 & \\ 4 & \end{array} \right)$

$$\text{is}(w) = 2, \quad \text{ds}(w) = 3.$$

# Erdős-Szekeres theorem

**Corollary** (Erdős-Szekeres, Seidenberg). *Let  $w \in \mathfrak{S}_{pq+1}$ . Then either  $\text{is}(w) > p$  or  $\text{ds}(w) > q$ .*



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**Proof.** Let  $\lambda = \text{sh}(w)$ . If  $\text{is}(w) \leq p$  and  $\text{ds}(w) \leq q$  then  $\lambda_1 \leq p$  and  $\lambda'_1 \leq q$ , so  $\sum \lambda_i \leq pq$ .  $\square$

# An extremal case

**Corollary.** *Say  $p \leq q$ . Then*

$$\begin{aligned} \#\{w \in \mathfrak{S}_{pq} : \text{is}(w) = p, \text{ds}(w) = q\} \\ = \left(f^{(p^q)}\right)^2 \end{aligned}$$

# An extremal case

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By hook-length formula, this is

$$\left( \frac{(pq)!}{1^1 2^2 \cdots p^p (p+1)^p \cdots q^p (q+1)^{p-1} \cdots (p+q-1)^1} \right)^2.$$

# Romik's theorem

**Romik:** let

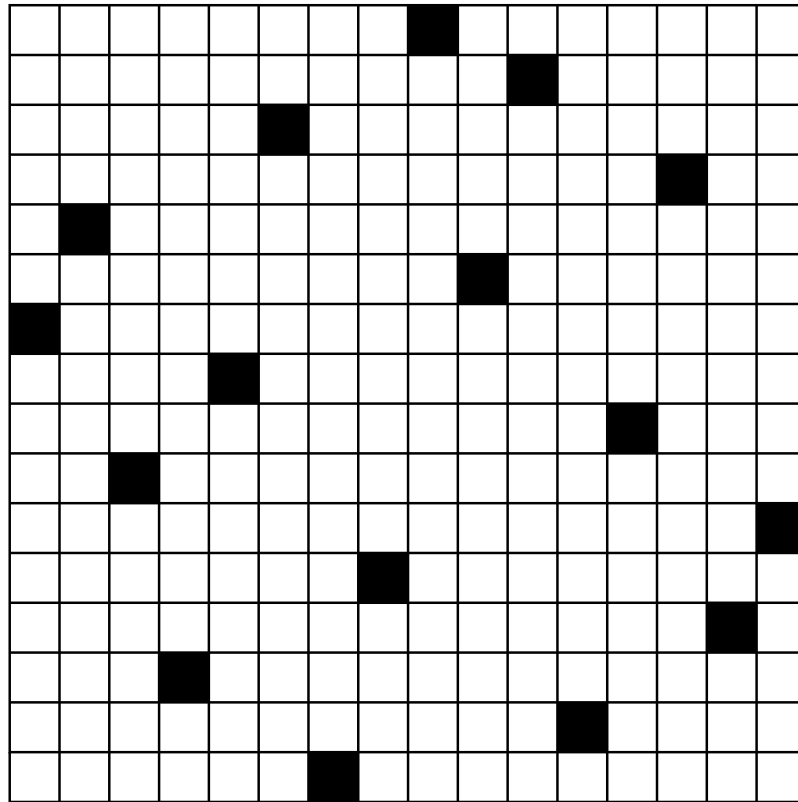
$$w \in \mathfrak{S}_{n^2}, \text{is}(w) = \text{ds}(w) = n.$$

Let  $P_w$  be the permutation matrix of  $w$  with corners  $(\pm 1, \pm 1)$ . Then (informally) as  $n \rightarrow \infty$  almost surely the 1's in  $P_w$  will become dense in the region bounded by the curve

$$(x^2 - y^2)^2 + 2(x^2 + y^2) = 3,$$

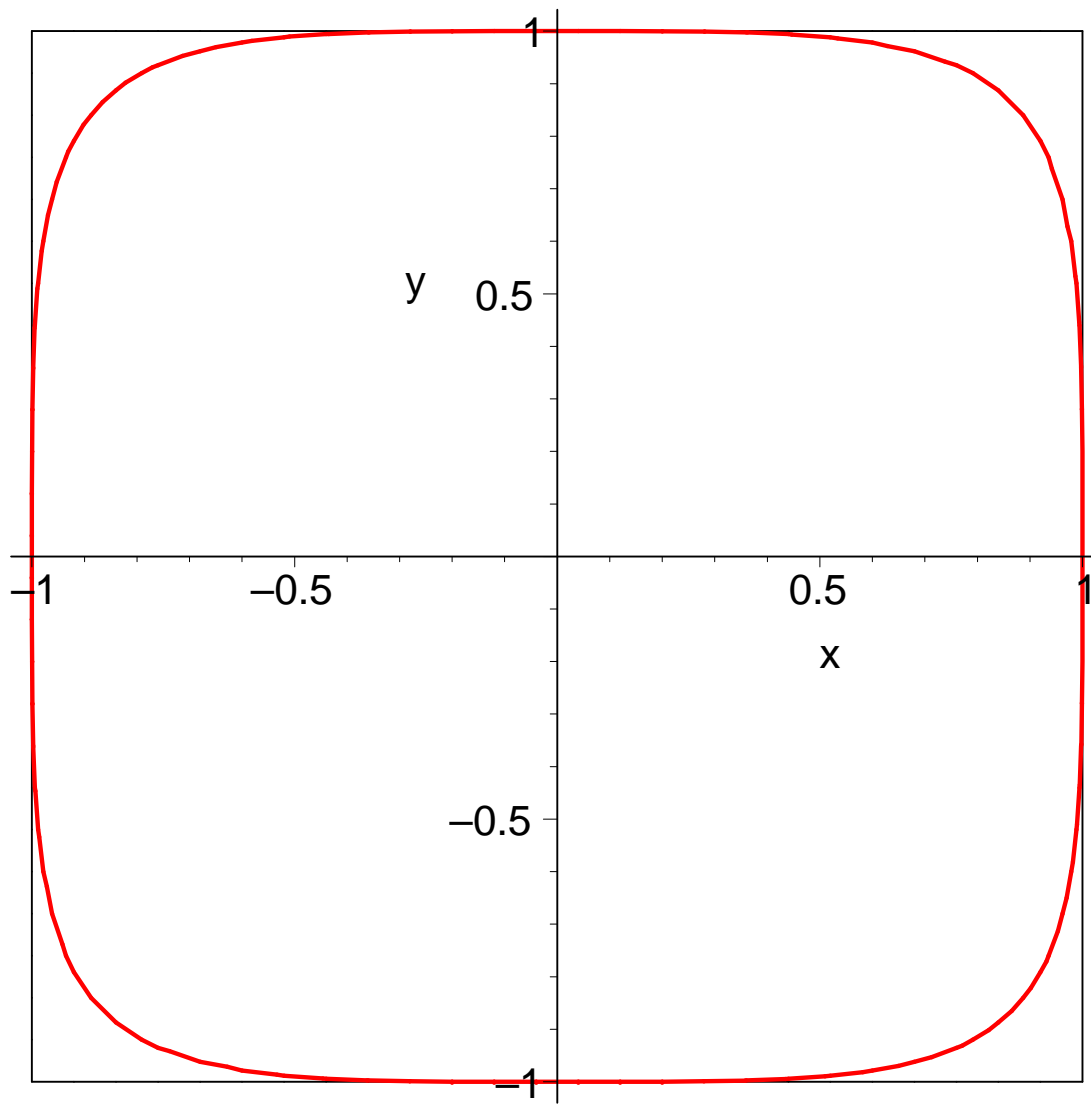
and will remain isolated outside this region.

# An example



$w = 9, 11, 6, 14, 2, 10, 1, 5, 13, 3, 16, 8, 15, 4, 12, 17$

$$(x^2 - y^2)^2 + 2(x^2 + y^2) = 3$$



# Area enclosed by curve

$$\begin{aligned}\alpha &= 8 \int_0^1 \frac{1}{\sqrt{(1-t^2)(1-(t/3)^2)}} dt \\ &\quad - 6 \int_0^1 \sqrt{\frac{1-(t/3)^2}{1-t^2}} dt \\ &= 4(0.94545962 \dots)\end{aligned}$$

# Expectation of $\text{is}(w)$

$$\begin{aligned} E(n) &= \text{expectation of } \text{is}(w), w \in \mathfrak{S}_n \\ &= \frac{1}{n!} \sum_{w \in \mathfrak{S}_n} \text{is}(w) \\ &= \frac{1}{n!} \sum_{\lambda \vdash n} \lambda_1 (f^\lambda)^2 \end{aligned}$$



# Expectation of $\text{is}(w)$

$E(n)$  = expectation of  $\text{is}(w)$ ,  $w \in \mathfrak{S}_n$

$$= \frac{1}{n!} \sum_{w \in \mathfrak{S}_n} \text{is}(w)$$

$$= \frac{1}{n!} \sum_{\lambda \vdash n} \lambda_1 (f^\lambda)^2$$

**Ulam:** what is distribution of  $\text{is}(w)$ ? rate of growth of  $E(n)$ ?

# Work of Hammersley

**Hammersley (1972):**

$$\exists c = \lim_{n \rightarrow \infty} n^{-1/2} E(n),$$

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Conjectured  $c = 2$ .

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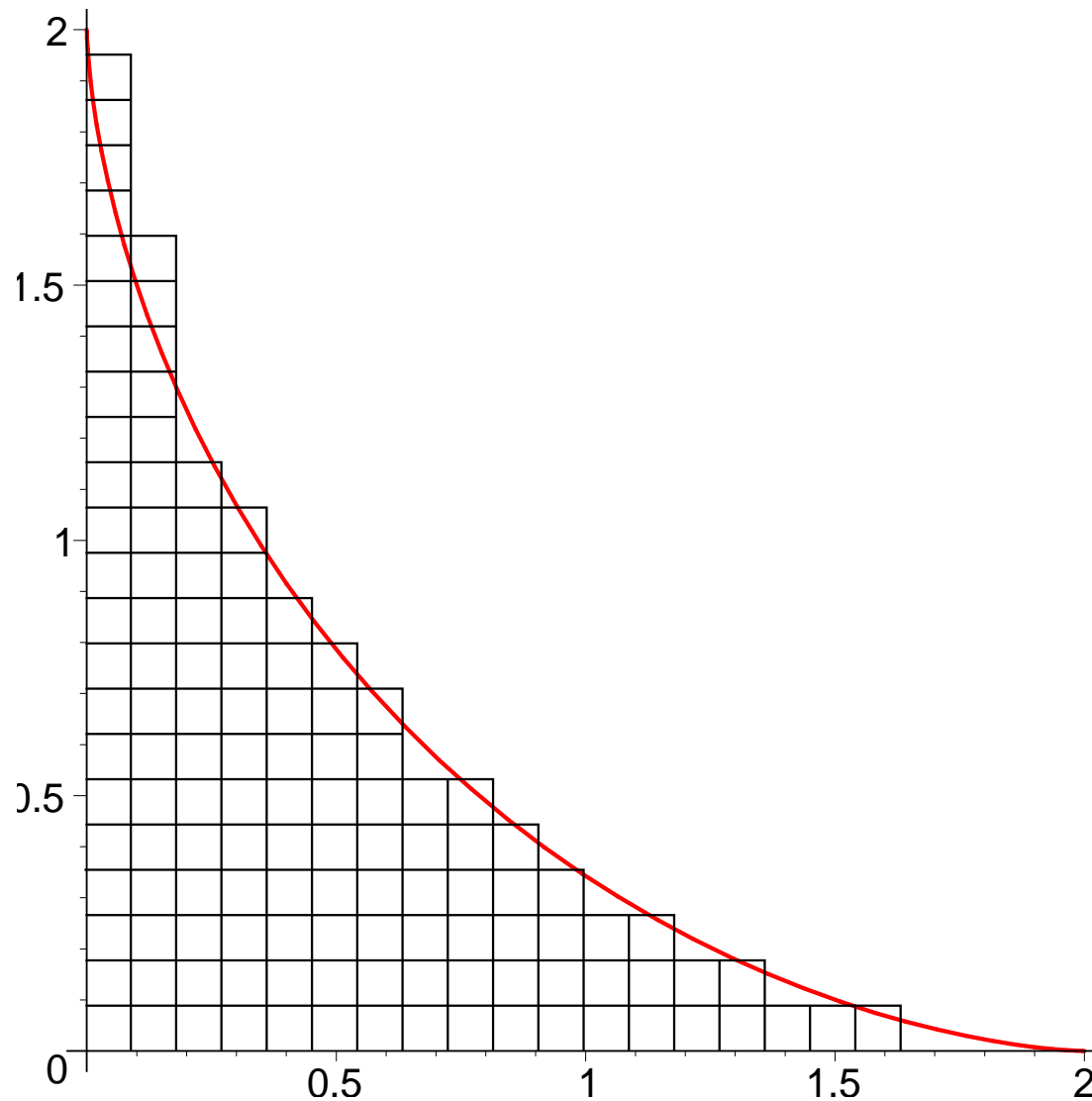
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**Idea of proof.**

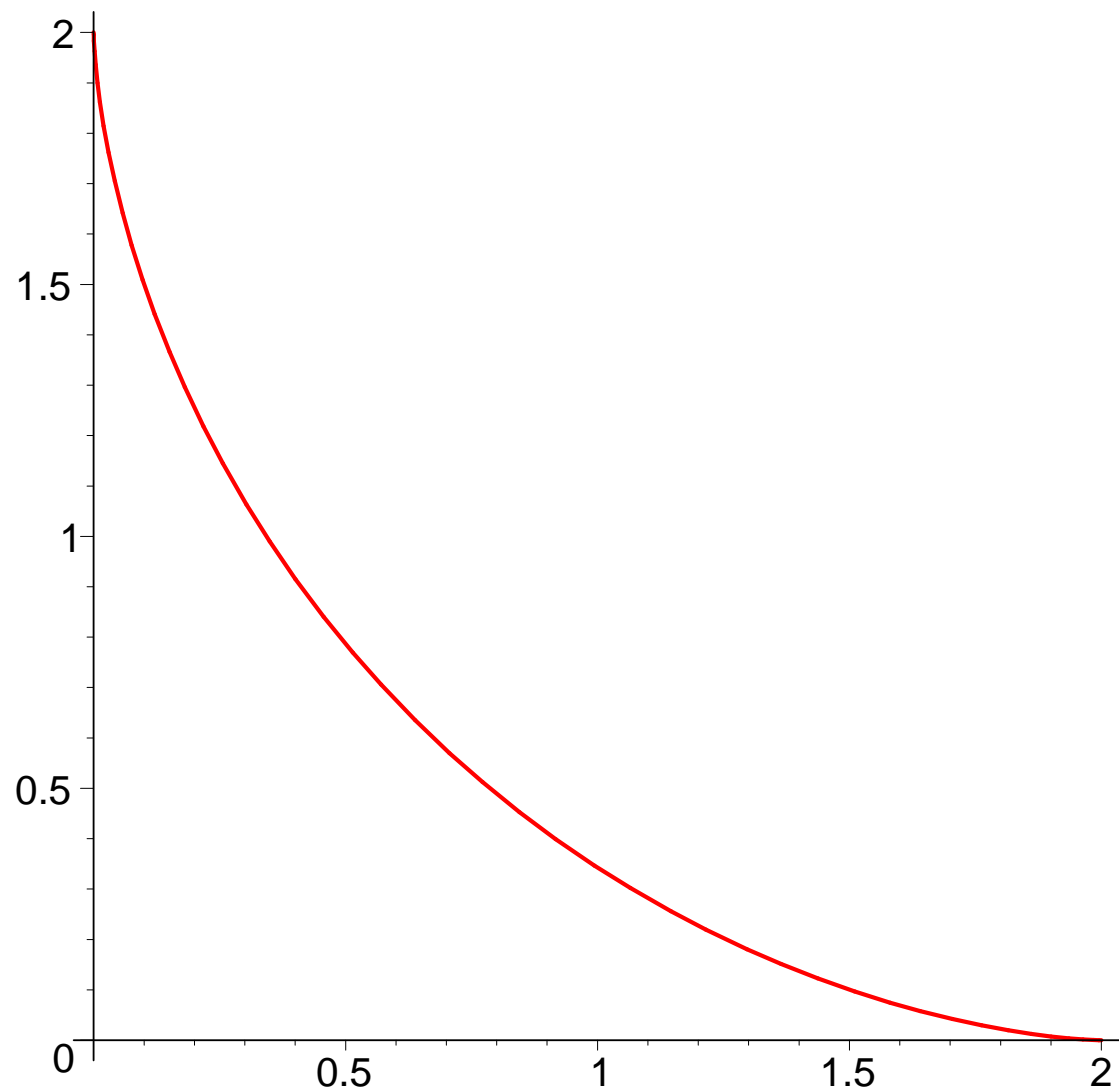
$$\begin{aligned} E(n) &= \frac{1}{n!} \sum_{\lambda \vdash n} \lambda_1 (f^\lambda)^2 \\ &\approx \frac{1}{n!} \max_{\lambda \vdash n} \lambda_1 (f^\lambda)^2. \end{aligned}$$

Find “limiting shape” of  $\lambda \vdash n$  maximizing  $\lambda$  as  $n \rightarrow \infty$  using hook-length formula.

# A big shape



# The limiting curve



# Equation of limiting curve

$$x = y + 2 \cos \theta$$

$$y = \frac{2}{\pi} (\sin \theta - \theta \cos \theta)$$

$$0 \leq \theta \leq \pi$$



$$\text{is}(w) \leq 2$$

$$u_k(n) := \#\{w \in \mathfrak{S}_n : \text{is}_n(w) \leq k\}.$$

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**J. M. Hammersley** (1972):

$$u_2(n) = C_n = \frac{1}{n+1} \binom{2n}{n},$$

a **Catalan number**.

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For  $\geq 160$  combinatorial interpretations of  $C_n$ , see

[www-math.mit.edu/~rstan/ec](http://www-math.mit.edu/~rstan/ec)

# Gessel's theorem

I. Gessel (1990):

$$\sum_{n \geq 0} u_k(n) \frac{x^{2n}}{n!^2} = \det [I_{|i-j|}(2x)]_{i,j=1}^k,$$

where

$$I_m(2x) = \sum_{j \geq 0} \frac{x^{m+2j}}{j!(m+j)!},$$

a **hyperbolic Bessel function** of the first kind of order  $m$ .

# The case $k = 2$

**Example.**  $\sum_{n \geq 0} u_2(n) \frac{x^{2n}}{n!^2}$

$$= I_0(2x)^2 - I_1(2x)^2$$

$$= \sum_{n \geq 0} C_n \frac{x^{2n}}{n!^2}.$$

# Painlevé II equation

**Baik-Deift-Johansson:**

Define  $u(x)$  by

$$\frac{d^2}{dx^2}u(x) = 2u(x)^3 + xu(x) \quad (*),$$

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$(*)$  is the **Painlevé II** equation (roughly, the branch points and essential singularities are independent of the initial conditions).

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**1933:** died in Paris.

# The Tracy-Widom distribution

$$F(t) = \exp \left( - \int_t^\infty (x - t) u(x)^2 dx \right)$$

where  $u(x)$  is the Painlevé II function.

# The Baik-Deift-Johansson theorem

Let  $\chi$  be a random variable with distribution  $F$ ,  
and let  $\chi_n$  be the random variable on  $\mathfrak{S}_n$ :

$$\chi_n(w) = \frac{iS_n(w) - 2\sqrt{n}}{n^{1/6}}.$$

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$$\chi_n(w) = \frac{\text{is}_n(w) - 2\sqrt{n}}{n^{1/6}}.$$

**Theorem.** As  $n \rightarrow \infty$ ,

$$\chi_n \rightarrow \chi \quad \text{in distribution,}$$

*i.e.*,

$$\lim_{n \rightarrow \infty} \text{Prob}(\chi_n \leq t) = F(t).$$

# Expectation redux

Recall  $E(n) \sim 2\sqrt{n}$ .



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**Corollary to BDJ theorem.**

$$\begin{aligned} E(n) &= 2\sqrt{n} + \left( \int t dF(t) \right) n^{1/6} + o(n^{1/6}) \\ &= 2\sqrt{n} - (1.7711 \dots) n^{1/6} + o(n^{1/6}) \end{aligned}$$

# Proof of BDJ theorem

Gessel's theorem reduces the problem to “just” analysis, viz., the **Riemann-Hilbert problem** in the theory of integrable systems, and the **method of steepest descent** to analyze the asymptotic behavior of integrable systems.

# Origin of Tracy-Widom distribution

Where did the Tracy-Widom distribution  $F(t)$  come from?

$$F(t) = \exp \left( - \int_t^\infty (x - t) u(x)^2 dx \right)$$

$$\frac{d^2}{dx^2} u(x) = 2u(x)^3 + xu(x)$$

# Gaussian Unitary Ensemble (GUE)

Analogue of normal distribution for  $n \times n$  hermitian matrices  $M = (M_{ij})$ :

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Analogue of normal distribution for  $n \times n$  hermitian matrices  $M = (M_{ij})$ :

$$Z_n^{-1} e^{-\text{tr}(M^2)} dM,$$

$$dM = \prod_i dM_{ii} \cdot \prod_{i < j} d(\Re M_{ij}) d(\Im M_{ij}),$$

where  $Z_n$  is a normalization constant.

# Tracy-Widom theorem

**Tracy-Widom** (1994): let  $\alpha_1$  denote the largest eigenvalue of  $M$ . Then

$$\lim_{n \rightarrow \infty} \text{Prob} \left( \left( \alpha_1 - \sqrt{2n} \right) \sqrt{2n}^{1/6} \leq t \right) = F(t).$$

# Random topologies

Is the connection between  $is(w)$  and GUE a coincidence?

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Okounkov provides a connection, via the theory of **random topologies on surfaces**. Very briefly, a surface can be described in two ways:

- Gluing polygons along their edges, connected to random matrices via quantum gravity.
- Ramified covering of a sphere, which can be formulated in terms of permutations.



# A variation

**Alternating sequence** of length  $k$ :

$$b_1 > b_2 < b_3 > b_4 < \cdots b_k$$

$E_n$ : number of alternating  $w \in \mathfrak{S}_n$  (**Euler number**)

$E_4 = 5$ : 2134, 3142, 3241, 4132, 4231

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$E_4 = 5$ : 2134, 3142, 3241, 4132, 4231

**Désiré André** (1840–1917): showed in 1879 that

$$\sum_{n \geq 0} E_n \frac{x^n}{n!} = \sec x + \tan x$$

# *Alternating subsequences?*

$as(w)$  = length of longest alternating subseq. of  $w$

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$$w = 56218347 \Rightarrow as(w) = 5$$

# The main lemma

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$$a_k(n) = \#\{w \in \mathfrak{S}_n : \text{as}(w) = k\}$$

$$\begin{aligned} b_k(n) &= a_1(n) + a_2(n) + \cdots + a_k(n) \\ &= \#\{w \in \mathfrak{S}_n : \text{as}(w) \leq k\}. \end{aligned}$$

# Recurrence for $a_k(n)$

$$\Rightarrow a_k(n) = \sum_{j=1}^n \binom{n-1}{j-1}$$

$$\sum_{2r+s=k-1} (a_{2r}(j-1) + a_{2r+1}(j-1)) a_s(n-j)$$

# $B(x, t)$ and $A(x, t)$

Define

$$B(x, t) = \sum_{k, n \geq 0} b_k(n) t^k \frac{x^n}{n!}$$

$$A(x, t) = \sum_{k, n \geq 0} a_k(n) t^k \frac{x^n}{n!}$$



# The main generating function

## Theorem.

$$B(x, t) = \frac{2/\rho}{1 - \frac{1-\rho}{t}e^{\rho x}} - \frac{1}{\rho}$$

$$A(x, t) = (1 - t)B(x, t),$$

where  $\rho = \sqrt{1 - t^2}$ .

# Formulas for $b_k(n)$

## Corollary.

$$\Rightarrow b_1(n) = 1$$

$$b_2(n) = n$$

$$b_3(n) = \frac{1}{4}(3^n - 2n + 3)$$

$$b_4(n) = \frac{1}{8}(4^n - (2n - 4)2^n)$$

⋮

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⋮

no such formulas for longest **increasing** subsequences

# Mean (expectation) of $as(w)$

$$D(n) = \frac{1}{n!} \sum_{w \in \mathfrak{S}_n} as(w),$$

the **expectation** of  $as(w)$  for  $w \in \mathfrak{S}_n$

# A formula for $D(n)$

$$\begin{aligned}\sum_{n \geq 1} D(n)x^n &= \frac{\partial}{\partial t} A(x, 1) \\ &= \frac{6x - 3x^2 + x^3}{6(1-x)^2} \\ &= x + \sum_{n \geq 2} \frac{4n+1}{6} x^n.\end{aligned}$$

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$$\Rightarrow D(n) = \frac{4n+1}{6}, \quad n \geq 2$$

# Comparison of $E(n)$ and $D(n)$

$$D(n) = \frac{4n + 1}{6}, \quad n \geq 2$$

$$E(n) \sim 2\sqrt{n}$$

# Variance of $as(w)$

$$V(n) = \frac{1}{n!} \sum_{w \in \mathfrak{S}_n} \left( as(w) - \frac{4n+1}{6} \right)^2, \quad n \geq 2$$

the **variance** of  $as(n)$  for  $w \in \mathfrak{S}_n$



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**Corollary.**

$$V(n) = \frac{8}{45}n - \frac{13}{180}, \quad n \geq 4$$

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the **variance** of  $as(w)$  for  $w \in \mathfrak{S}_n$

**Corollary.**

$$V(n) = \frac{8}{45}n - \frac{13}{180}, \quad n \geq 4$$

similar results for higher moments

# A new distribution?

$$P(t) = \lim_{n \rightarrow \infty} \text{Prob}_{w \in \mathfrak{S}_n} \left( \frac{\text{as}_n(w) - 2n/3}{\sqrt{n}} \leq t \right)$$

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Stanley distribution?

# Limiting distribution

**Theorem** (Pemantle, Widom, (Wilf)).

$$\lim_{n \rightarrow \infty} \text{Prob}_{w \in \mathfrak{S}_n} \left( \frac{\text{as}(w) - 2n/3}{\sqrt{n}} \leq t \right)$$

$$= \frac{1}{\sqrt{\pi}} \int_{-\infty}^{t\sqrt{45}/4} e^{-s^2} ds$$

(Gaussian distribution)

# Limiting distribution

**Theorem** (Pemantle, Widom, (Wilf)).

$$\lim_{n \rightarrow \infty} \text{Prob}_{w \in \mathfrak{S}_n} \left( \frac{\text{as}(w) - 2n/3}{\sqrt{n}} \leq t \right)$$

$$= \frac{1}{\sqrt{\pi}} \int_{-\infty}^{t\sqrt{45}/4} e^{-s^2} ds$$

(Gaussian distribution)



# $k$ -alternating sequences

Given  $k \geq 1$ , define a sequence  $a_1 a_2 \cdots a_n$  of integers to be  **$k$ -alternating** if

$$a_i > a_{i+1} \Leftrightarrow i \equiv 1 \pmod{k}.$$

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**Example.** 61482572 is 3-alternating



# $a_k(w)$ and $E_k(n)$

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$$a_{n-1}(w) = \text{is}(w)$$

$$a_2(w) = \text{as}(w)$$

$$E_k(n) = \text{expectation of } a_k(w)$$

$$= \frac{1}{n!} \sum_{w \in \mathfrak{S}_n} a_k(w)$$

# A problem

$E_k(n)$  interpolates between  $E(n) \sim 2\sqrt{n}$  and  $D(n) \sim 2n/3$ . Is there a sharp cutoff between  $c\sqrt{n}$  and  $cn$  behavior, or do we get intermediate values like  $cn^\alpha$ ,  $\frac{1}{2} < \alpha < 1$ , say for  $k = \sqrt{n}$ ?

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Similar questions for the limiting distribution: do we interpolate between Tracy-Widom and Gaussian?

# A variant

Same questions if we replace  $k$ -alternating with:

$$a_i > a_{i+1} \Leftrightarrow \lfloor i/k \rfloor \text{ is even.}$$

E.g.,  $k = 3$ :

$$a_1 > a_2 > a_3 < a_4 < a_5 > a_6 > a_7 < \dots$$

