#### 18.600: Lecture 23

## Sums of independent random variables

Scott Sheffield

MIT

18.600 Lecture 23

Say we have independent random variables X and Y and we know their density functions f<sub>X</sub> and f<sub>Y</sub>.

- Say we have independent random variables X and Y and we know their density functions f<sub>X</sub> and f<sub>Y</sub>.
- Now let's try to find  $F_{X+Y}(a) = P\{X + Y \le a\}$ .

- Say we have independent random variables X and Y and we know their density functions f<sub>X</sub> and f<sub>Y</sub>.
- Now let's try to find  $F_{X+Y}(a) = P\{X + Y \le a\}$ .
- ► This is the integral over {(x, y) : x + y ≤ a} of f(x, y) = f<sub>X</sub>(x)f<sub>Y</sub>(y). Thus,

- Say we have independent random variables X and Y and we know their density functions f<sub>X</sub> and f<sub>Y</sub>.
- Now let's try to find  $F_{X+Y}(a) = P\{X + Y \le a\}$ .
- This is the integral over  $\{(x, y) : x + y \le a\}$  of  $f(x, y) = f_X(x)f_Y(y)$ . Thus,

$$P\{X + Y \le a\} = \int_{-\infty}^{\infty} \int_{-\infty}^{a-y} f_X(x) f_Y(y) dx dy$$
$$= \int_{-\infty}^{\infty} F_X(a-y) f_Y(y) dy.$$

- Say we have independent random variables X and Y and we know their density functions f<sub>X</sub> and f<sub>Y</sub>.
- Now let's try to find  $F_{X+Y}(a) = P\{X + Y \le a\}$ .
- ► This is the integral over  $\{(x, y) : x + y \le a\}$  of  $f(x, y) = f_X(x)f_Y(y)$ . Thus,

$$P\{X + Y \le a\} = \int_{-\infty}^{\infty} \int_{-\infty}^{a-y} f_X(x) f_Y(y) dx dy$$
$$= \int_{-\infty}^{\infty} F_X(a-y) f_Y(y) dy.$$

• Differentiating both sides gives  $f_{X+Y}(a) = \frac{d}{da} \int_{-\infty}^{\infty} F_X(a-y) f_Y(y) dy = \int_{-\infty}^{\infty} f_X(a-y) f_Y(y) dy.$ 

- Say we have independent random variables X and Y and we know their density functions f<sub>X</sub> and f<sub>Y</sub>.
- Now let's try to find  $F_{X+Y}(a) = P\{X + Y \le a\}$ .
- This is the integral over  $\{(x, y) : x + y \le a\}$  of  $f(x, y) = f_X(x)f_Y(y)$ . Thus,

$$P\{X + Y \le a\} = \int_{-\infty}^{\infty} \int_{-\infty}^{a-y} f_X(x) f_Y(y) dx dy$$
$$= \int_{-\infty}^{\infty} F_X(a-y) f_Y(y) dy.$$

- Differentiating both sides gives  $f_{X+Y}(a) = \frac{d}{da} \int_{-\infty}^{\infty} F_X(a-y) f_Y(y) dy = \int_{-\infty}^{\infty} f_X(a-y) f_Y(y) dy.$
- Latter formula makes some intuitive sense. We're integrating over the set of x, y pairs that add up to a.

18.600 Lecture 23

The abbreviation i.i.d. means independent identically distributed.

- The abbreviation i.i.d. means independent identically distributed.
- It is actually one of the most important abbreviations in probability theory.

- The abbreviation i.i.d. means independent identically distributed.
- It is actually one of the most important abbreviations in probability theory.
- Worth memorizing.

Suppose that X and Y are i.i.d. and uniform on [0,1]. So  $f_X = f_Y = 1$  on [0,1].

- Suppose that X and Y are i.i.d. and uniform on [0, 1]. So  $f_X = f_Y = 1$  on [0, 1].
- What is the probability density function of X + Y?

- Suppose that X and Y are i.i.d. and uniform on [0,1]. So  $f_X = f_Y = 1$  on [0,1].
- ▶ What is the probability density function of *X* + *Y*?
- ►  $f_{X+Y}(a) = \int_{-\infty}^{\infty} f_X(a-y) f_Y(y) dy = \int_0^1 f_X(a-y)$  which is the length of  $[0,1] \cap [a-1,a]$ .

- Suppose that X and Y are i.i.d. and uniform on [0,1]. So  $f_X = f_Y = 1$  on [0,1].
- What is the probability density function of X + Y?
- ►  $f_{X+Y}(a) = \int_{-\infty}^{\infty} f_X(a-y) f_Y(y) dy = \int_0^1 f_X(a-y)$  which is the length of  $[0,1] \cap [a-1,a]$ .
- ► That's a when a ∈ [0, 1] and 2 − a when a ∈ [1, 2] and 0 otherwise.

## Review: summing i.i.d. geometric random variables

• A geometric random variable X with parameter p has  $P\{X = k\} = (1 - p)^{k-1}p$  for  $k \ge 1$ .

# Review: summing i.i.d. geometric random variables

- A geometric random variable X with parameter p has  $P\{X = k\} = (1 p)^{k-1}p$  for  $k \ge 1$ .
- Sum Z of n independent copies of X?

- A geometric random variable X with parameter p has  $P\{X = k\} = (1 p)^{k-1}p$  for  $k \ge 1$ .
- Sum Z of n independent copies of X?
- We can interpret Z as time slot where nth head occurs in i.i.d. sequence of p-coin tosses.

- A geometric random variable X with parameter p has  $P\{X = k\} = (1 p)^{k-1}p$  for  $k \ge 1$ .
- Sum Z of n independent copies of X?
- We can interpret Z as time slot where nth head occurs in i.i.d. sequence of p-coin tosses.
- ► So Z is negative binomial (n, p). So  $P\{Z = k\} = {\binom{k-1}{n-1}}p^{n-1}(1-p)^{k-n}p$ .

Suppose  $X_1, \ldots X_n$  are i.i.d. exponential random variables with parameter  $\lambda$ . So  $f_{X_i}(x) = \lambda e^{-\lambda x}$  on  $[0, \infty)$  for all  $1 \le i \le n$ .

▶ Suppose  $X_1, ..., X_n$  are i.i.d. exponential random variables with parameter  $\lambda$ . So  $f_{X_i}(x) = \lambda e^{-\lambda x}$  on  $[0, \infty)$  for all  $1 \le i \le n$ .

• What is the law of 
$$Z = \sum_{i=1}^{n} X_i$$
?

- ▶ Suppose  $X_1, ..., X_n$  are i.i.d. exponential random variables with parameter  $\lambda$ . So  $f_{X_i}(x) = \lambda e^{-\lambda x}$  on  $[0, \infty)$  for all  $1 \le i \le n$ .
- What is the law of  $Z = \sum_{i=1}^{n} X_i$ ?
- We claimed in an earlier lecture that this was a gamma distribution with parameters (λ, n).

- ▶ Suppose  $X_1, ..., X_n$  are i.i.d. exponential random variables with parameter  $\lambda$ . So  $f_{X_i}(x) = \lambda e^{-\lambda x}$  on  $[0, \infty)$  for all  $1 \le i \le n$ .
- What is the law of  $Z = \sum_{i=1}^{n} X_i$ ?
- We claimed in an earlier lecture that this was a gamma distribution with parameters (λ, n).

• So 
$$f_Z(y) = \frac{\lambda e^{-\lambda y} (\lambda y)^{n-1}}{\Gamma(n)}$$
.

- ▶ Suppose  $X_1, ..., X_n$  are i.i.d. exponential random variables with parameter  $\lambda$ . So  $f_{X_i}(x) = \lambda e^{-\lambda x}$  on  $[0, \infty)$  for all  $1 \le i \le n$ .
- What is the law of  $Z = \sum_{i=1}^{n} X_i$ ?
- We claimed in an earlier lecture that this was a gamma distribution with parameters (λ, n).

• So 
$$f_Z(y) = \frac{\lambda e^{-\lambda y} (\lambda y)^{n-1}}{\Gamma(n)}$$
.

We argued this point by taking limits of negative binomial distributions. Can we check it directly?

- ▶ Suppose  $X_1, ..., X_n$  are i.i.d. exponential random variables with parameter  $\lambda$ . So  $f_{X_i}(x) = \lambda e^{-\lambda x}$  on  $[0, \infty)$  for all  $1 \le i \le n$ .
- What is the law of  $Z = \sum_{i=1}^{n} X_i$ ?
- We claimed in an earlier lecture that this was a gamma distribution with parameters (λ, n).

• So 
$$f_Z(y) = \frac{\lambda e^{-\lambda y} (\lambda y)^{n-1}}{\Gamma(n)}$$
.

- We argued this point by taking limits of negative binomial distributions. Can we check it directly?
- By induction, would suffice to show that a gamma (λ, 1) plus an independent gamma (λ, n) is a gamma (λ, n + 1).

Say X is gamma (λ, s), Y is gamma (λ, t), and X and Y are independent.

- Say X is gamma (λ, s), Y is gamma (λ, t), and X and Y are independent.
- Intuitively, X is amount of time till we see s events, and Y is amount of subsequent time till we see t more events.

- Say X is gamma (λ, s), Y is gamma (λ, t), and X and Y are independent.
- Intuitively, X is amount of time till we see s events, and Y is amount of subsequent time till we see t more events.

• So 
$$f_X(x) = \frac{\lambda e^{-\lambda x} (\lambda x)^{s-1}}{\Gamma(s)}$$
 and  $f_Y(y) = \frac{\lambda e^{-\lambda y} (\lambda y)^{t-1}}{\Gamma(t)}$ .

- Say X is gamma (λ, s), Y is gamma (λ, t), and X and Y are independent.
- Intuitively, X is amount of time till we see s events, and Y is amount of subsequent time till we see t more events.

• So 
$$f_X(x) = \frac{\lambda e^{-\lambda x} (\lambda x)^{s-1}}{\Gamma(s)}$$
 and  $f_Y(y) = \frac{\lambda e^{-\lambda y} (\lambda y)^{t-1}}{\Gamma(t)}$ .

• Now 
$$f_{X+Y}(a) = \int_{-\infty}^{\infty} f_X(a-y)f_Y(y)dy$$
.

- Say X is gamma (λ, s), Y is gamma (λ, t), and X and Y are independent.
- Intuitively, X is amount of time till we see s events, and Y is amount of subsequent time till we see t more events.

• So 
$$f_X(x) = \frac{\lambda e^{-\lambda x} (\lambda x)^{s-1}}{\Gamma(s)}$$
 and  $f_Y(y) = \frac{\lambda e^{-\lambda y} (\lambda y)^{t-1}}{\Gamma(t)}$ .

• Now 
$$f_{X+Y}(a) = \int_{-\infty}^{\infty} f_X(a-y) f_Y(y) dy$$
.

Up to an a-independent multiplicative constant, this is

$$\int_0^a e^{-\lambda(a-y)} (a-y)^{s-1} e^{-\lambda y} y^{t-1} dy = e^{-\lambda a} \int_0^a (a-y)^{s-1} y^{t-1} dy.$$

- Say X is gamma (λ, s), Y is gamma (λ, t), and X and Y are independent.
- Intuitively, X is amount of time till we see s events, and Y is amount of subsequent time till we see t more events.

• So 
$$f_X(x) = \frac{\lambda e^{-\lambda x} (\lambda x)^{s-1}}{\Gamma(s)}$$
 and  $f_Y(y) = \frac{\lambda e^{-\lambda y} (\lambda y)^{t-1}}{\Gamma(t)}$ .

• Now 
$$f_{X+Y}(a) = \int_{-\infty}^{\infty} f_X(a-y) f_Y(y) dy$$
.

Up to an a-independent multiplicative constant, this is

$$\int_0^a e^{-\lambda(a-y)} (a-y)^{s-1} e^{-\lambda y} y^{t-1} dy = e^{-\lambda a} \int_0^a (a-y)^{s-1} y^{t-1} dy.$$

• Letting 
$$x = y/a$$
, this becomes  
 $e^{-\lambda a}a^{s+t-1}\int_0^1 (1-x)^{s-1}x^{t-1}dx$ .

- Say X is gamma (λ, s), Y is gamma (λ, t), and X and Y are independent.
- Intuitively, X is amount of time till we see s events, and Y is amount of subsequent time till we see t more events.

• So 
$$f_X(x) = \frac{\lambda e^{-\lambda x} (\lambda x)^{s-1}}{\Gamma(s)}$$
 and  $f_Y(y) = \frac{\lambda e^{-\lambda y} (\lambda y)^{t-1}}{\Gamma(t)}$ .

• Now 
$$f_{X+Y}(a) = \int_{-\infty}^{\infty} f_X(a-y)f_Y(y)dy$$
.

Up to an a-independent multiplicative constant, this is

$$\int_0^a e^{-\lambda(a-y)} (a-y)^{s-1} e^{-\lambda y} y^{t-1} dy = e^{-\lambda a} \int_0^a (a-y)^{s-1} y^{t-1} dy.$$

- Letting x = y/a, this becomes  $e^{-\lambda a}a^{s+t-1}\int_0^1 (1-x)^{s-1}x^{t-1}dx$ .
- This is (up to multiplicative constant) e<sup>-λa</sup>a<sup>s+t-1</sup>. Constant must be such that integral from −∞ to ∞ is 1. Conclude that X + Y is gamma (λ, s + t).

18.600 Lecture 23

X is normal with mean zero, variance σ<sup>2</sup><sub>1</sub>, Y is normal with mean zero, variance σ<sup>2</sup><sub>2</sub>.

X is normal with mean zero, variance σ<sup>2</sup><sub>1</sub>, Y is normal with mean zero, variance σ<sup>2</sup><sub>2</sub>.

• 
$$f_X(x) = \frac{1}{\sqrt{2\pi\sigma_1}} e^{\frac{-x^2}{2\sigma_1^2}}$$
 and  $f_Y(y) = \frac{1}{\sqrt{2\pi\sigma_2}} e^{\frac{-y^2}{2\sigma_2^2}}$ .

X is normal with mean zero, variance σ<sup>2</sup><sub>1</sub>, Y is normal with mean zero, variance σ<sup>2</sup><sub>2</sub>.

• 
$$f_X(x) = \frac{1}{\sqrt{2\pi\sigma_1}} e^{\frac{-x^2}{2\sigma_1^2}}$$
 and  $f_Y(y) = \frac{1}{\sqrt{2\pi\sigma_2}} e^{\frac{-y^2}{2\sigma_2^2}}$ .

• We just need to compute  $f_{X+Y}(a) = \int_{-\infty}^{\infty} f_X(a-y) f_Y(y) dy$ .

X is normal with mean zero, variance σ<sup>2</sup><sub>1</sub>, Y is normal with mean zero, variance σ<sup>2</sup><sub>2</sub>.

• 
$$f_X(x) = \frac{1}{\sqrt{2\pi\sigma_1}} e^{\frac{-x^2}{2\sigma_1^2}}$$
 and  $f_Y(y) = \frac{1}{\sqrt{2\pi\sigma_2}} e^{\frac{-y^2}{2\sigma_2^2}}$ .

• We just need to compute  $f_{X+Y}(a) = \int_{-\infty}^{\infty} f_X(a-y) f_Y(y) dy$ .

• We could compute this directly.

X is normal with mean zero, variance σ<sup>2</sup><sub>1</sub>, Y is normal with mean zero, variance σ<sup>2</sup><sub>2</sub>.

• 
$$f_X(x) = \frac{1}{\sqrt{2\pi\sigma_1}} e^{\frac{-x^2}{2\sigma_1^2}}$$
 and  $f_Y(y) = \frac{1}{\sqrt{2\pi\sigma_2}} e^{\frac{-y^2}{2\sigma_2^2}}$ .

- We just need to compute  $f_{X+Y}(a) = \int_{-\infty}^{\infty} f_X(a-y) f_Y(y) dy$ .
- We could compute this directly.
- Or we could argue with a multi-dimensional bell curve picture that if X and Y have variance 1 then f<sub>σ1X+σ2Y</sub> is the density of a normal random variable (and note that variances and expectations are additive).

X is normal with mean zero, variance σ<sup>2</sup><sub>1</sub>, Y is normal with mean zero, variance σ<sup>2</sup><sub>2</sub>.

• 
$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma_1} e^{\frac{-x^2}{2\sigma_1^2}}$$
 and  $f_Y(y) = \frac{1}{\sqrt{2\pi}\sigma_2} e^{\frac{-y^2}{2\sigma_2^2}}$ .

- We just need to compute  $f_{X+Y}(a) = \int_{-\infty}^{\infty} f_X(a-y) f_Y(y) dy$ .
- We could compute this directly.
- Or we could argue with a multi-dimensional bell curve picture that if X and Y have variance 1 then f<sub>σ1X+σ2Y</sub> is the density of a normal random variable (and note that variances and expectations are additive).
- Or use fact that if  $A_i \in \{-1, 1\}$  are i.i.d. coin tosses then  $\frac{1}{\sqrt{N}} \sum_{i=1}^{\sigma^2 N} A_i$  is approximately normal with variance  $\sigma^2$  when N is large.

X is normal with mean zero, variance σ<sup>2</sup><sub>1</sub>, Y is normal with mean zero, variance σ<sup>2</sup><sub>2</sub>.

• 
$$f_X(x) = \frac{1}{\sqrt{2\pi\sigma_1}} e^{\frac{-x^2}{2\sigma_1^2}}$$
 and  $f_Y(y) = \frac{1}{\sqrt{2\pi\sigma_2}} e^{\frac{-y^2}{2\sigma_2^2}}$ .

- We just need to compute  $f_{X+Y}(a) = \int_{-\infty}^{\infty} f_X(a-y) f_Y(y) dy$ .
- We could compute this directly.
- Or we could argue with a multi-dimensional bell curve picture that if X and Y have variance 1 then f<sub>σ1X+σ2Y</sub> is the density of a normal random variable (and note that variances and expectations are additive).
- Or use fact that if  $A_i \in \{-1, 1\}$  are i.i.d. coin tosses then  $\frac{1}{\sqrt{N}} \sum_{i=1}^{\sigma^2 N} A_i$  is approximately normal with variance  $\sigma^2$  when N is large.
- Generally: if independent random variables  $X_j$  are normal  $(\mu_j, \sigma_j^2)$  then  $\sum_{j=1}^n X_j$  is normal  $(\sum_{j=1}^n \mu_j, \sum_{j=1}^n \sigma_j^2)$ .

18.600 Lecture 23

#### ▶ Sum of an independent binomial (*m*, *p*) and binomial (*n*, *p*)?

- ▶ Sum of an independent binomial (*m*, *p*) and binomial (*n*, *p*)?
- ► Yes, binomial (m + n, p). Can be seen from coin toss interpretation.

- ▶ Sum of an independent binomial (*m*, *p*) and binomial (*n*, *p*)?
- ► Yes, binomial (m + n, p). Can be seen from coin toss interpretation.
- Sum of independent Poisson  $\lambda_1$  and Poisson  $\lambda_2$ ?

- Sum of an independent binomial (m, p) and binomial (n, p)?
- ► Yes, binomial (m + n, p). Can be seen from coin toss interpretation.
- Sum of independent Poisson  $\lambda_1$  and Poisson  $\lambda_2$ ?
- ► Yes, Poisson λ<sub>1</sub> + λ<sub>2</sub>. Can be seen from Poisson point process interpretation.