# 18.600: Lecture 13 Lectures 1-12 Review

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#### Counting tricks and basic principles of probability

Discrete random variables

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Discrete random variables

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- ► Answer: <sup>(n+k-1)</sup>
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   Represent partition by k - 1 bars and n stars, e.g., as \*\* | \*\*|| \*\*\*\*|\*.

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- Countable additivity:  $P(\bigcup_{i=1}^{\infty} E_i) = \sum_{i=1}^{\infty} P(E_i)$  if  $E_i \cap E_j = \emptyset$  for each pair *i* and *j*.

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- More generally,

$$P(\cup_{i=1}^{n} E_{i}) = \sum_{i=1}^{n} P(E_{i}) - \sum_{i_{1} < i_{2}} P(E_{i_{1}} E_{i_{2}}) + \dots + (-1)^{(r+1)} \sum_{i_{1} < i_{2} < \dots < i_{r}} P(E_{i_{1}} E_{i_{2}} \dots E_{i_{r}})$$
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► The notation ∑<sub>i1 < i2 < ... < ir</sub> means a sum over all of the <sup>n</sup>/<sub>r</sub> subsets of size r of the set {1, 2, ..., n}.

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$$P(\bigcup_{i=1}^{n} E_i) = 1 - \frac{1}{2!} + \frac{1}{3!} - \frac{1}{4!} + \ldots \pm \frac{1}{n!}$$

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- ►  $1 P(\bigcup_{i=1}^{n} E_i) = 1 1 + \frac{1}{2!} \frac{1}{3!} + \frac{1}{4!} \ldots \pm \frac{1}{n!} \approx 1/e \approx .36788$

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- ► Nice fact:  $P(E_1E_2E_3...E_n) = P(E_1)P(E_2|E_1)P(E_3|E_1E_2)...P(E_n|E_1...E_{n-1})$

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- Useful when we think about multi-step experiments.
- For example, let E<sub>i</sub> be event ith person gets own hat in the n-hat shuffle problem.

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$$P(E) = P(EF) + P(EF^{c})$$
  
=  $P(E|F)P(F) + P(E|F^{c})P(F^{c})$ 

Dividing probability into two cases

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In words: want to know the probability of *E*. There are two scenarios *F* and *F<sup>c</sup>*. If I know the probabilities of the two scenarios and the probability of *E* conditioned on each scenario, I can work out the probability of *E*.

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• So 
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▶ Ratio 
$$\frac{P(B|A)}{P(B)}$$
 determines "how compelling new evidence is".

We can check the probability axioms: 0 ≤ P(E|F) ≤ 1, P(S|F) = 1, and P(∪E<sub>i</sub>) = ∑ P(E<sub>i</sub>|F), if *i* ranges over a countable set and the E<sub>i</sub> are disjoint. We can check the probability axioms: 0 ≤ P(E|F) ≤ 1, P(S|F) = 1, and P(∪E<sub>i</sub>) = ∑ P(E<sub>i</sub>|F), if *i* ranges over a countable set and the E<sub>i</sub> are disjoint.

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- The probability measure  $P(\cdot|F)$  is related to  $P(\cdot)$ .
- ► To get former from latter, we set probabilities of elements outside of F to zero and multiply probabilities of events inside of F by 1/P(F).
- ► P(·) is the prior probability measure and P(·|F) is the posterior measure (revised after discovering that F occurs).

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- Say *E* and *F* are **independent** if P(EF) = P(E)P(F).
- ► Equivalent statement: P(E|F) = P(E). Also equivalent: P(F|E) = P(F).

► Say  $E_1 \dots E_n$  are independent if for each  $\{i_1, i_2, \dots, i_k\} \subset \{1, 2, \dots, n\}$  we have  $P(E_{i_1}E_{i_2} \dots E_{i_k}) = P(E_{i_1})P(E_{i_2}) \dots P(E_{i_k}).$ 

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- ▶ Independence implies  $P(E_1E_2E_3|E_4E_5E_6) = \frac{P(E_1)P(E_2)P(E_3)P(E_4)P(E_5)P(E_6)}{P(E_4)P(E_5)P(E_6)} = P(E_1E_2E_3)$ , and other similar statements.

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- No. Consider these three events: first coin heads, second coin heads, odd number heads. Pairwise independent, not independent.

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- ► For each a in this countable set, write p(a) := P{X = a}. Call p the probability mass function.
- Write F(a) = P{X ≤ a} = ∑<sub>x≤a</sub> p(x). Call F the cumulative distribution function.

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- Example: in *n*-hat shuffle problem, let *E<sub>i</sub>* be the event *i*th person gets own hat.
- Then  $\sum_{i=1}^{n} 1_{E_i}$  is total number of people who get own hats.

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Represents weighted average of possible values X can take, each value being weighted by its probability. If the state space S is countable, we can give SUM OVER
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Agrees with the SUM OVER POSSIBLE X VALUES definition:

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- In fact, for real constants a and b, we have E[aX + bY] = aE[X] + bE[Y].
- This is called the **linearity of expectation**.
- Can extend to more variables  $E[X_1 + X_2 + \ldots + X_n] = E[X_1] + E[X_2] + \ldots + E[X_n].$

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- The variance of X, denoted Var(X), is defined by Var(X) = E[(X − μ)<sup>2</sup>].
- ► Taking  $g(x) = (x \mu)^2$ , and recalling that  $E[g(X)] = \sum_{x:p(x)>0} g(x)p(x)$ , we find that

$$\operatorname{Var}[X] = \sum_{x: p(x) > 0} (x - \mu)^2 p(x).$$

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- The variance of X, denoted Var(X), is defined by Var(X) = E[(X − μ)<sup>2</sup>].
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- Very important alternate formula:  $Var[X] = E[X^2] (E[X])^2$ .

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• Proof:  $\operatorname{Var}[aX] = E[a^2X^2] - E[aX]^2 = a^2E[X^2] - a^2E[X]^2 = a^2\operatorname{Var}[X].$ 

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- If we switch from feet to inches in our "height of randomly chosen person" example, then X, E[X], and SD[X] each get multiplied by 12, but Var[X] gets multiplied by 144.

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- Note that  $E[X_j] = p \cdot 1 + (1-p) \cdot 0 = p$  for each j.
- Conclude by additivity of expectation that

$$E[X] = \sum_{j=1}^{n} E[X_j] = \sum_{j=1}^{n} p = np.$$

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• Can show generally that if  $X_1, \ldots, X_n$  independent then  $\operatorname{Var}[\sum_{j=1}^n X_j] = \sum_{j=1}^n \operatorname{Var}[X_j]$ 

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- This also suggests  $E[X] = np = \lambda$  and  $Var[X] = npq \approx \lambda$ .

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- Probability to see zero events in first t time units is  $e^{-\lambda t}$ .
- Let  $T_k$  be time elapsed, since the previous event, until the *k*th event occurs. Then the  $T_k$  are independent random variables, each of which is exponential with parameter  $\lambda$ .

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