18.440: Lecture 17 Continuous random variables

Scott Sheffield

MIT

Outline

Outline

Say X is a continuous random variable if there exists a probability density function $f = f_X$ on \mathbb{R} such that $P\{X \in B\} = \int_B f(x) dx := \int \mathbb{1}_B(x) f(x) dx$.

▲□▶ ▲□▶ ▲目▶ ▲目▶ - 目 - のへで

- Say X is a continuous random variable if there exists a probability density function $f = f_X$ on \mathbb{R} such that $P\{X \in B\} = \int_B f(x) dx := \int 1_B(x) f(x) dx$.
- We may assume ∫_ℝ f(x)dx = ∫_{-∞}[∞] f(x)dx = 1 and f is non-negative.

- Say X is a continuous random variable if there exists a probability density function $f = f_X$ on \mathbb{R} such that $P\{X \in B\} = \int_B f(x) dx := \int 1_B(x) f(x) dx$.
- We may assume ∫_ℝ f(x)dx = ∫_{-∞}[∞] f(x)dx = 1 and f is non-negative.
- Probability of interval [a, b] is given by ∫_a^b f(x)dx, the area under f between a and b.

- Say X is a continuous random variable if there exists a probability density function $f = f_X$ on \mathbb{R} such that $P\{X \in B\} = \int_B f(x) dx := \int 1_B(x) f(x) dx$.
- We may assume ∫_ℝ f(x)dx = ∫_{-∞}[∞] f(x)dx = 1 and f is non-negative.
- Probability of interval [a, b] is given by ∫_a^b f(x)dx, the area under f between a and b.

Probability of any single point is zero.

- ▶ Say X is a continuous random variable if there exists a probability density function $f = f_X$ on \mathbb{R} such that $P\{X \in B\} = \int_B f(x)dx := \int \mathbb{1}_B(x)f(x)dx$.
- We may assume ∫_ℝ f(x)dx = ∫_{-∞}[∞] f(x)dx = 1 and f is non-negative.
- Probability of interval [a, b] is given by ∫_a^b f(x)dx, the area under f between a and b.
- Probability of any single point is zero.
- Define cumulative distribution function $F(a) = F_X(a) := P\{X < a\} = P\{X \le a\} = \int_{-\infty}^a f(x) dx.$

► Suppose
$$f(x) = \begin{cases} 1/2 & x \in [0,2] \\ 0 & x \notin [0,2]. \end{cases}$$

Suppose
$$f(x) = \begin{cases} 1/2 & x \in [0, 2] \\ 0 & x \notin [0, 2]. \end{cases}$$
What is $P\{X < 3/2\}$?

◆□▶ ◆□▶ ◆三▶ ◆三▶ 三三 - のへで

▶ Suppose f(x) =

$$\begin{cases}
1/2 & x \in [0, 2] \\
0 & x \notin [0, 2].
\end{cases}$$

▶ What is P{X < 3/2}?

◆□▶ ◆□▶ ◆三▶ ◆三▶ 三三 - のへで

► Suppose
$$f(x) = \begin{cases} 1/2 & x \in [0,2] \\ 0 & x \notin [0,2]. \end{cases}$$

◆□▶ ◆□▶ ◆臣▶ ◆臣▶ 臣 の�?

▶ What is P{X < 3/2}?</p>

• What is
$$P\{1/2 < X < 3/2\}$$
?

► Suppose
$$f(x) = \begin{cases} 1/2 & x \in [0,2] \\ 0 & x \notin [0,2]. \end{cases}$$

▶ What is P{X < 3/2}?</p>

- ▶ What is P{1/2 < X < 3/2}?</p>
- What is $P\{X \in (0,1) \cup (3/2,5)\}$?

► Suppose
$$f(x) = \begin{cases} 1/2 & x \in [0,2] \\ 0 & x \notin [0,2]. \end{cases}$$

▶ What is P{X < 3/2}?</p>

- What is $P\{1/2 < X < 3/2\}$?
- What is $P\{X \in (0,1) \cup (3/2,5)\}$?

What is F?

► Suppose
$$f(x) = \begin{cases} 1/2 & x \in [0,2] \\ 0 & x \notin [0,2]. \end{cases}$$

▶ What is P{X < 3/2}?</p>

- What is $P\{1/2 < X < 3/2\}$?
- What is $P\{X \in (0,1) \cup (3/2,5)\}$?
- What is F?
- ► We say that X is uniformly distributed on the interval [0, 2].

▲□▶ ▲□▶ ▲臣▶ ▲臣▶ □ □ ● ●

► Suppose
$$f(x) = \begin{cases} x/2 & x \in [0,2] \\ 0 & 0 \notin [0,2]. \end{cases}$$

◆□▶ ◆□▶ ◆臣▶ ◆臣▶ □臣 = のへで

▶ Suppose
$$f(x) = \begin{cases} x/2 & x \in [0,2] \\ 0 & 0 \notin [0,2]. \end{cases}$$
▶ What is $P\{X < 3/2\}$?

◆□▶ ◆□▶ ◆臣▶ ◆臣▶ □臣 = のへで

▶ Suppose
$$f(x) = \begin{cases} x/2 & x \in [0,2] \\ 0 & 0 \notin [0,2]. \end{cases}$$
▶ What is $P\{X < 3/2\}?$

• What is
$$P\{X = 3/2\}$$
?

► Suppose
$$f(x) = \begin{cases} x/2 & x \in [0,2] \\ 0 & 0 \notin [0,2]. \end{cases}$$

- ▶ What is P{X = 3/2}?
- What is $P\{1/2 < X < 3/2\}$?

► Suppose
$$f(x) = \begin{cases} x/2 & x \in [0,2] \\ 0 & 0 \notin [0,2]. \end{cases}$$

- ▶ What is P{X = 3/2}?
- What is $P\{1/2 < X < 3/2\}$?
- ▶ What is *F*?

Outline

Outline

► Recall that when X was a discrete random variable, with p(x) = P{X = x}, we wrote

$$E[X] = \sum_{x:p(x)>0} p(x)x.$$

▲□▶ ▲□▶ ▲臣▶ ★臣▶ 三臣 - のへで

► Recall that when X was a discrete random variable, with p(x) = P{X = x}, we wrote

$$E[X] = \sum_{x:p(x)>0} p(x)x.$$

▲□▶ ▲□▶ ▲□▶ ▲□▶ ▲□ ● ● ●

► How should we define E[X] when X is a continuous random variable?

► Recall that when X was a discrete random variable, with p(x) = P{X = x}, we wrote

$$E[X] = \sum_{x:p(x)>0} p(x)x.$$

- ► How should we define E[X] when X is a continuous random variable?
- Answer: $E[X] = \int_{-\infty}^{\infty} f(x) x dx$.

► Recall that when X was a discrete random variable, with p(x) = P{X = x}, we wrote

$$E[X] = \sum_{x:p(x)>0} p(x)x.$$

- ► How should we define E[X] when X is a continuous random variable?
- Answer: $E[X] = \int_{-\infty}^{\infty} f(x) x dx$.
- ► Recall that when X was a discrete random variable, with p(x) = P{X = x}, we wrote

$$E[g(X)] = \sum_{x:p(x)>0} p(x)g(x).$$

▲□▶ ▲□▶ ▲臣▶ ★臣▶ 三臣 - のへで

► Recall that when X was a discrete random variable, with p(x) = P{X = x}, we wrote

$$E[X] = \sum_{x:p(x)>0} p(x)x.$$

- ► How should we define E[X] when X is a continuous random variable?
- Answer: $E[X] = \int_{-\infty}^{\infty} f(x) x dx$.
- ► Recall that when X was a discrete random variable, with p(x) = P{X = x}, we wrote

$$E[g(X)] = \sum_{x:p(x)>0} p(x)g(x).$$

What is the analog when X is a continuous random variable?

► Recall that when X was a discrete random variable, with p(x) = P{X = x}, we wrote

$$E[X] = \sum_{x:p(x)>0} p(x)x.$$

- ► How should we define E[X] when X is a continuous random variable?
- Answer: $E[X] = \int_{-\infty}^{\infty} f(x) x dx$.
- ► Recall that when X was a discrete random variable, with p(x) = P{X = x}, we wrote

$$E[g(X)] = \sum_{x:p(x)>0} p(x)g(x).$$

- What is the analog when X is a continuous random variable?
- Answer: we will write $E[g(X)] = \int_{-\infty}^{\infty} f(x)g(x)dx$.

18.440 Lecture 17

Suppose X is a continuous random variable with mean μ .

◆□▶ ◆□▶ ◆臣▶ ◆臣▶ 臣 の�?

- Suppose X is a continuous random variable with mean μ .
- We can write Var[X] = E[(X − µ)²], same as in the discrete case.

- Suppose X is a continuous random variable with mean μ .
- We can write Var[X] = E[(X − µ)²], same as in the discrete case.

• Next, if
$$g = g_1 + g_2$$
 then
 $E[g(X)] = \int g_1(x)f(x)dx + \int g_2(x)f(x)dx = \int (g_1(x) + g_2(x))f(x)dx = E[g_1(X)] + E[g_2(X)].$

- Suppose X is a continuous random variable with mean μ .
- We can write Var[X] = E[(X − µ)²], same as in the discrete case.

• Next, if
$$g = g_1 + g_2$$
 then
 $E[g(X)] = \int g_1(x)f(x)dx + \int g_2(x)f(x)dx = \int (g_1(x) + g_2(x))f(x)dx = E[g_1(X)] + E[g_2(X)].$

Furthermore, E[ag(X)] = aE[g(X)] when a is a constant.

▲□▶ ▲□▶ ▲臣▶ ★臣▶ 三臣 - のへで

- Suppose X is a continuous random variable with mean μ .
- We can write Var[X] = E[(X − µ)²], same as in the discrete case.

• Next, if
$$g = g_1 + g_2$$
 then
 $E[g(X)] = \int g_1(x)f(x)dx + \int g_2(x)f(x)dx = \int (g_1(x) + g_2(x))f(x)dx = E[g_1(X)] + E[g_2(X)].$

- Furthermore, E[ag(X)] = aE[g(X)] when a is a constant.
- Just as in the discrete case, we can expand the variance expression as Var[X] = E[X² − 2µX + µ²] and use additivity of expectation to say that Var[X] = E[X²] − 2µE[X] + E[µ²] = E[X²] − 2µ² + µ² = E[X²] − E[X]².

- Suppose X is a continuous random variable with mean μ .
- We can write Var[X] = E[(X − µ)²], same as in the discrete case.

• Next, if
$$g = g_1 + g_2$$
 then
 $E[g(X)] = \int g_1(x)f(x)dx + \int g_2(x)f(x)dx = \int (g_1(x) + g_2(x))f(x)dx = E[g_1(X)] + E[g_2(X)].$

- Furthermore, E[ag(X)] = aE[g(X)] when a is a constant.
- ► Just as in the discrete case, we can expand the variance expression as $Var[X] = E[X^2 2\mu X + \mu^2]$ and use additivity of expectation to say that $Var[X] = E[X^2] 2\mu E[X] + E[\mu^2] = E[X^2] 2\mu^2 + \mu^2 = E[X^2] E[X]^2$.
- This formula is often useful for calculations.

• Suppose that
$$f_X(x) = \begin{cases} 1/2 & x \in [0,2] \\ 0 & x \notin [0,2]. \end{cases}$$

◆□▶ ◆□▶ ◆臣▶ ◆臣▶ 臣 の�?

Suppose that
$$f_X(x) = \begin{cases} 1/2 & x \in [0,2] \\ 0 & x \notin [0,2]. \end{cases}$$

◆□▶ ◆□▶ ◆臣▶ ◆臣▶ 臣 の�?

▶ What is Var[X]?

• Suppose that
$$f_X(x) = \begin{cases} 1/2 & x \in [0,2] \\ 0 & x \notin [0,2]. \end{cases}$$

▶ What is Var[X]?

• Suppose instead that
$$f_X(x) = \begin{cases} x/2 & x \in [0,2] \\ 0 & 0 \notin [0,2]. \end{cases}$$

◆□▶ ◆□▶ ◆臣▶ ◆臣▶ 臣 の�?

• Suppose that
$$f_X(x) = \begin{cases} 1/2 & x \in [0,2] \\ 0 & x \notin [0,2]. \end{cases}$$

▶ What is Var[X]?

• Suppose instead that
$$f_X(x) = \begin{cases} x/2 & x \in [0,2] \\ 0 & 0 \notin [0,2]. \end{cases}$$

▶ What is Var[X]?

Outline

Outline

• One of the very simplest probability density functions is $f(x) = \begin{cases} 1 & x \in [0, 1] \\ 0 & 0 \notin [0, 1]. \end{cases}$

- One of the very simplest probability density functions is $f(x) = \begin{cases} 1 & x \in [0, 1] \\ 0 & 0 \notin [0, 1]. \end{cases}$
- If B ⊂ [0, 1] is an interval, then P{X ∈ B} is the length of that interval.

- One of the very simplest probability density functions is $f(x) = \begin{cases} 1 & x \in [0, 1] \\ 0 & 0 \notin [0, 1]. \end{cases}$
- If B ⊂ [0,1] is an interval, then P{X ∈ B} is the length of that interval.
- ▶ Generally, if $B \subset [0,1]$ then $P\{X \in B\} = \int_B 1 dx = \int 1_B(x) dx$ is the "total volume" or "total length" of the set *B*.

- One of the very simplest probability density functions is $f(x) = \begin{cases} 1 & x \in [0, 1] \\ 0 & 0 \notin [0, 1]. \end{cases}$
- If B ⊂ [0,1] is an interval, then P{X ∈ B} is the length of that interval.
- Generally, if $B \subset [0,1]$ then $P\{X \in B\} = \int_B 1 dx = \int 1_B(x) dx$ is the "total volume" or "total length" of the set B.

▲□▶ ▲□▶ ▲□▶ ▲□▶ ▲□ ● ● ●

▶ What if *B* is the set of all rational numbers?

- One of the very simplest probability density functions is $f(x) = \begin{cases} 1 & x \in [0, 1] \\ 0 & 0 \notin [0, 1]. \end{cases}$
- If B ⊂ [0,1] is an interval, then P{X ∈ B} is the length of that interval.
- Generally, if $B \subset [0,1]$ then $P\{X \in B\} = \int_B 1 dx = \int 1_B(x) dx$ is the "total volume" or "total length" of the set B.
- What if *B* is the set of all rational numbers?
- ► How do we mathematically define the volume of an arbitrary set B?

 Uniform probability measure on [0, 1) should satisfy translation invariance: If B and a horizontal translation of B are both subsets [0, 1), their probabilities should be equal.

- Uniform probability measure on [0, 1) should satisfy translation invariance: If B and a horizontal translation of B are both subsets [0, 1), their probabilities should be equal.
- Consider wrap-around translations $\tau_r(x) = (x + r) \mod 1$.

- Uniform probability measure on [0, 1) should satisfy translation invariance: If B and a horizontal translation of B are both subsets [0, 1), their probabilities should be equal.
- Consider wrap-around translations $\tau_r(x) = (x + r) \mod 1$.

▲□▶ ▲□▶ ▲□▶ ▲□▶ ▲□ ● ● ●

• By translation invariance, $\tau_r(B)$ has same probability as B.

- Uniform probability measure on [0, 1) should satisfy translation invariance: If B and a horizontal translation of B are both subsets [0, 1), their probabilities should be equal.
- Consider wrap-around translations $\tau_r(x) = (x + r) \mod 1$.
- By translation invariance, $\tau_r(B)$ has same probability as B.
- Call x, y "equivalent modulo rationals" if x − y is rational (e.g., x = π − 3 and y = π − 9/4). An equivalence class is the set of points in [0, 1) equivalent to some given point.

オロト 本理 トイヨト オヨト ヨー ろくつ

- Uniform probability measure on [0, 1) should satisfy translation invariance: If B and a horizontal translation of B are both subsets [0, 1), their probabilities should be equal.
- Consider wrap-around translations $\tau_r(x) = (x + r) \mod 1$.
- By translation invariance, $\tau_r(B)$ has same probability as B.
- Call x, y "equivalent modulo rationals" if x − y is rational (e.g., x = π − 3 and y = π − 9/4). An equivalence class is the set of points in [0, 1) equivalent to some given point.

There are uncountably many of these classes.

- Uniform probability measure on [0, 1) should satisfy translation invariance: If B and a horizontal translation of B are both subsets [0, 1), their probabilities should be equal.
- Consider wrap-around translations $\tau_r(x) = (x + r) \mod 1$.
- By translation invariance, $\tau_r(B)$ has same probability as B.
- Call x, y "equivalent modulo rationals" if x − y is rational (e.g., x = π − 3 and y = π − 9/4). An equivalence class is the set of points in [0, 1) equivalent to some given point.
- There are uncountably many of these classes.
- Let A ⊂ [0, 1) contain one point from each class. For each x ∈ [0, 1), there is one a ∈ A such that r = x − a is rational.

- Uniform probability measure on [0, 1) should satisfy translation invariance: If B and a horizontal translation of B are both subsets [0, 1), their probabilities should be equal.
- Consider wrap-around translations $\tau_r(x) = (x + r) \mod 1$.
- By translation invariance, $\tau_r(B)$ has same probability as B.
- Call x, y "equivalent modulo rationals" if x − y is rational (e.g., x = π − 3 and y = π − 9/4). An equivalence class is the set of points in [0, 1) equivalent to some given point.
- There are uncountably many of these classes.
- Let A ⊂ [0, 1) contain one point from each class. For each x ∈ [0, 1), there is one a ∈ A such that r = x − a is rational.
- ▶ Then each x in [0,1) lies in $\tau_r(A)$ for **one** rational $r \in [0,1)$.

- Uniform probability measure on [0, 1) should satisfy translation invariance: If B and a horizontal translation of B are both subsets [0, 1), their probabilities should be equal.
- Consider wrap-around translations $\tau_r(x) = (x + r) \mod 1$.
- By translation invariance, $\tau_r(B)$ has same probability as B.
- Call x, y "equivalent modulo rationals" if x − y is rational (e.g., x = π − 3 and y = π − 9/4). An equivalence class is the set of points in [0, 1) equivalent to some given point.
- There are uncountably many of these classes.
- Let A ⊂ [0, 1) contain one point from each class. For each x ∈ [0, 1), there is one a ∈ A such that r = x − a is rational.
- ▶ Then each x in [0,1) lies in $\tau_r(A)$ for **one** rational $r \in [0,1)$.
- Thus $[0,1) = \cup \tau_r(A)$ as r ranges over rationals in [0,1).

- Uniform probability measure on [0, 1) should satisfy translation invariance: If B and a horizontal translation of B are both subsets [0, 1), their probabilities should be equal.
- Consider wrap-around translations $\tau_r(x) = (x + r) \mod 1$.
- By translation invariance, $\tau_r(B)$ has same probability as B.
- Call x, y "equivalent modulo rationals" if x − y is rational (e.g., x = π − 3 and y = π − 9/4). An equivalence class is the set of points in [0, 1) equivalent to some given point.
- There are uncountably many of these classes.
- Let A ⊂ [0, 1) contain one point from each class. For each x ∈ [0, 1), there is one a ∈ A such that r = x − a is rational.
- ▶ Then each x in [0,1) lies in $\tau_r(A)$ for **one** rational $r \in [0,1)$.
- Thus $[0,1) = \cup \tau_r(A)$ as r ranges over rationals in [0,1).
- ▶ If P(A) = 0, then $P(S) = \sum_{r} P(\tau_r(A)) = 0$. If P(A) > 0 then $P(S) = \sum_{r} P(\tau_r(A)) = \infty$. Contradicts P(S) = 1 axiom.

18.440 Lecture 17

1. Re-examine axioms of mathematics: the very existence of a set A with one element from each equivalence class is consequence of so-called axiom of choice. Removing that axiom makes paradox goes away, since one can just suppose (pretend?) these kinds of sets don't exist.

イロン 不通 と 不通 と 不通 と 一道

- ► 1. Re-examine axioms of mathematics: the very existence of a set A with one element from each equivalence class is consequence of so-called axiom of choice. Removing that axiom makes paradox goes away, since one can just suppose (pretend?) these kinds of sets don't exist.
- 2. Re-examine axioms of probability: Replace countable additivity with finite additivity? (Look up Banach-Tarski.)

イロン 不通 と 不通 と 不通 と 一道

- ► 1. Re-examine axioms of mathematics: the very existence of a set A with one element from each equivalence class is consequence of so-called axiom of choice. Removing that axiom makes paradox goes away, since one can just suppose (pretend?) these kinds of sets don't exist.
- 2. Re-examine axioms of probability: Replace countable additivity with finite additivity? (Look up Banach-Tarski.)
- 3. Keep the axiom of choice and countable additivity but don't define probabilities of all sets: Instead of defining P(B) for every subset B of sample space, restrict attention to a family of so-called "measurable" sets.

- ► 1. Re-examine axioms of mathematics: the very existence of a set A with one element from each equivalence class is consequence of so-called axiom of choice. Removing that axiom makes paradox goes away, since one can just suppose (pretend?) these kinds of sets don't exist.
- 2. Re-examine axioms of probability: Replace countable additivity with finite additivity? (Look up Banach-Tarski.)
- 3. Keep the axiom of choice and countable additivity but don't define probabilities of all sets: Instead of defining P(B) for every subset B of sample space, restrict attention to a family of so-called "measurable" sets.
- Most mainstream probability and analysis takes the third approach.

- ► 1. Re-examine axioms of mathematics: the very existence of a set A with one element from each equivalence class is consequence of so-called axiom of choice. Removing that axiom makes paradox goes away, since one can just suppose (pretend?) these kinds of sets don't exist.
- 2. Re-examine axioms of probability: Replace countable additivity with finite additivity? (Look up Banach-Tarski.)
- 3. Keep the axiom of choice and countable additivity but don't define probabilities of all sets: Instead of defining P(B) for every subset B of sample space, restrict attention to a family of so-called "measurable" sets.
- Most mainstream probability and analysis takes the third approach.
- In practice, sets we care about (e.g., countable unions of points and intervals) tend to be measurable.

18.440 Lecture 17

More advanced courses in probability and analysis (such as 18.125 and 18.175) spend a significant amount of time rigorously constructing a class of so-called **measurable sets** and the so-called **Lebesgue measure**, which assigns a real number (a measure) to each of these sets.

- More advanced courses in probability and analysis (such as 18.125 and 18.175) spend a significant amount of time rigorously constructing a class of so-called **measurable sets** and the so-called **Lebesgue measure**, which assigns a real number (a measure) to each of these sets.
- These courses also replace the Riemann integral with the so-called Lebesgue integral.

◆□▶ ◆□▶ ◆三▶ ◆三▶ □ のへ⊙

- More advanced courses in probability and analysis (such as 18.125 and 18.175) spend a significant amount of time rigorously constructing a class of so-called **measurable sets** and the so-called **Lebesgue measure**, which assigns a real number (a measure) to each of these sets.
- These courses also replace the Riemann integral with the so-called Lebesgue integral.

◆□▶ ◆□▶ ◆三▶ ◆三▶ □ のへ⊙

• We will not treat these topics any further in this course.

- More advanced courses in probability and analysis (such as 18.125 and 18.175) spend a significant amount of time rigorously constructing a class of so-called **measurable sets** and the so-called **Lebesgue measure**, which assigns a real number (a measure) to each of these sets.
- These courses also replace the Riemann integral with the so-called Lebesgue integral.
- We will not treat these topics any further in this course.
- We usually limit our attention to probability density functions f and sets B for which the ordinary Riemann integral ∫ 1_B(x)f(x)dx is well defined.

- More advanced courses in probability and analysis (such as 18.125 and 18.175) spend a significant amount of time rigorously constructing a class of so-called **measurable sets** and the so-called **Lebesgue measure**, which assigns a real number (a measure) to each of these sets.
- These courses also replace the Riemann integral with the so-called Lebesgue integral.
- We will not treat these topics any further in this course.
- We usually limit our attention to probability density functions f and sets B for which the ordinary Riemann integral ∫ 1_B(x)f(x)dx is well defined.
- Riemann integration is a mathematically rigorous theory. It's just not as robust as Lebesgue integration.