1 Introduction.

Let $x_i$ be a probability distribution on a totally ordered set $I$, i.e., the probability of $i \in I$ is $x_i$. (Hence $x_i \geq 0$ and $\sum x_i = 1$.) Fix $n \in \mathbb{P} = \{1, 2, \ldots\}$, and define a random permutation $w \in \mathfrak{S}_n$ as follows. For each $1 \leq j \leq n$, choose independently an integer $i_j$ (from the distribution $x_i$). Then standardize the sequence $i = i_1 \cdots i_n$ in the sense of [34, p. 322], i.e., let $\alpha_1 < \cdots < \alpha_k$ be the elements of $I$ actually appearing in $i$, and let $a_i$ be the number of $\alpha_i$’s in $i$. Replace the $\alpha_1$’s in $i$ by $1, 2, \ldots, a_1$ from left-to-right, then the $\alpha_2$’s in $i$ by $a_1 + 1, a_1 + 2, \ldots, a_1 + a_2$ from left-to-right, etc. For instance, if $I = \mathbb{P}$ and $i = 311431$, then $w = 412653$. This defines a probability distribution on the symmetric group $\mathfrak{S}_n$, which we call the $QS$-distribution (because of the close connection with quasisymmetric functions explained below). If we need to be explicit about the parameters $x = (x_i)_{i \in I}$, then we will refer to the $QS(x)$-distribution.

The QS-distribution is not new, at least when $I$ is finite. It appears for instance in [10, pp. 153–154][14][20][24]. Although these authors anticipate some of our results, they don’t make systematic use of quasisymmetric and symmetric functions. Some additional work using the viewpoint of our paper is due to Fulman [15].

1.1 Example. As an example of a general result to be proved later (Theorem 2.1), let us compute $\text{Prob}(213)$, the probability that a random permutation $w \in \mathfrak{S}_n$ has 213 as a pattern.

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permutation \( w \in \mathfrak{S}_3 \) (chosen from the QS-distribution) is equal to the permutation 213. A sequence \( \mathbf{i} = i_1i_2i_3 \) will have the standardization 213 if and only if \( i_2 < i_1 \leq i_3 \). Hence

\[
\text{Prob}(213) = \sum_{a < b \leq c} x_a x_b x_c.
\]

There is an alternative way to describe the QS-distribution. Suppose we have a deck of \( n \) cards. Cut the deck into random packets of respective sizes \( a_i > 0 , i \in I \), such that \( \sum a_i = n \) and the probability of \( (a_i)_{i \in I} \) is

\[
\text{Prob} ((a_i)_{i \in I}) = n! \prod_{i \in I} \frac{x_i^{a_i}}{a_i!}.
\]

Then riffle shuffle these packets \( P_i \) together into a pile \( P \) in the manner described by Bayer and Diaconis [2] (see also [11][14]); namely, after placing \( k \) cards on \( P \), the probability that the next card comes from the top of the current packet \( P_j \) is proportional to the current number of cards in \( P_j \). This card is then placed at the bottom of \( P \). The ordinary dovetail or riffle shuffle [2, p. 294] corresponds to the case \( I = \{1,2\} \) and \( x_1 = x_2 = 1/2 \). In this case the original deck is cut according to the binomial distribution. More generally, if for fixed \( q \in \mathbb{P} \) we cut the deck into \( q \) packets (some possibly empty) according to the \( q \)-multinomial distribution, then we obtain the \( q \)-shuffles of Bayer and Diaconis [2, p. 299] (where they use \( a \) for our \( q \)). The \( q \)-shuffle is identical to the QS-distribution for \( I = \{1,2 \ldots ,q\} \) and \( x_1 = x_2 = \cdots = x_q = 1/q \). We will denote this distribution by \( U_q \). If we \( q \)-shuffle a deck and then \( r \)-shuffle it, the distribution is the same as a single \( qr \)-shuffle [2, Lemma 1]. In other words, if \( * \) denotes convolution of probability distributions then \( U_q * U_r = U_{qr} \). We extend this result to the QS-distribution in Theorem 2.4.

In the next section we will establish the connection between the QS-distribution and the theory of quasisymmetric functions. In Section 3 we use known results from the theory of quasisymmetric and symmetric functions to obtain results about the QS-distribution, many of them direct generalizations of the work of Bayer, Diaconis, and Fulman mentioned above. For instance, we show that the probability that a random permutation \( w \in \mathfrak{S}_n \) (always assumed chosen from the QS-distribution) has \( k \) inversions is equal to the probability that \( w \) has major index \( k \). This is an analogue of MacMahon's
corresponding result for the uniform distribution on $\mathcal{S}_n$ (see e.g. [33, p. 23]). A further result is that if $T'$ is a standard Young tableau of shape $\lambda$, then the probability that $T'$ is the recording tableau of $w$ under the Robinson-Schensted-Knuth algorithm is $s_\lambda(x)$, where $x = (x_i)_{i \in I}$ and $s_\lambda$ denotes a Schur function.

2 The QS-distribution and quasisymmetric functions

In this section we show the connection between the QS-distribution and quasisymmetric functions. In general our reference for quasisymmetric functions and symmetric functions will be [34, Ch. 7]. Quasisymmetric functions may be regarded as bearing the same relation to compositions (ordered partitions of a nonnegative integer) as symmetric functions bear to partitions. Namely, the homogeneous symmetric functions of degree $n$ form a vector space of dimension $p(n)$ (the number of partitions of $n$) and have many natural bases indexed by the partitions of $n$. Similarly, the homogeneous quasisymmetric functions of degree $n$ form a vector space of dimension $2^n - 1$ (the number of compositions of $n$) and have many natural bases indexed by the compositions of $n$.

A quasisymmetric function $F(z)$ is a power series of bounded degree, say with rational coefficients, in the variables $z = (z_i)_{i \in I}$, with the following property: let $i_1 < i_2 < \cdots < i_n$ and $j_1 < j_2 < \cdots < j_n$, where $i_k, j_k \in I$. If $a_1, \ldots, a_k \in \mathbb{P}$, then the coefficient of $z_{i_1}^{a_1} \cdots z_{i_n}^{a_n}$ is equal to the coefficient of $z_{j_1}^{a_1} \cdots z_{j_n}^{a_n}$ in $F(z)$. The set of all quasisymmetric functions forms a graded $\mathbb{Q}$-algebra denoted $\mathbb{Q}$. The quasisymmetric functions that are homogeneous of degree $n$ form a $\mathbb{Q}$-vector space, denoted $\mathbb{Q}_n$. If $|I| \geq n$ (including $|I| = \infty$), then $\dim \mathbb{Q}_n = 2^{n-1}$.

Let $\alpha = (\alpha_1, \ldots, \alpha_k)$ be a composition of $n$, i.e., $\alpha_i \in \mathbb{P}$ and $\sum \alpha_i = n$. Let Comp$(n)$ denote the set of all compositions of $n$, so $\#\text{Comp}(n) = 2^{n-1}$ [33, pp. 14–15]. Define

$$S_\alpha = \{\alpha_1, \alpha_1 + \alpha_2, \ldots, \alpha_1 + \cdots + \alpha_{k-1}\} \subseteq \{1, \ldots, n-1\}.$$
The fundamental quasisymmetric function $L_\alpha = L_\alpha(z)$ is defined by

$$L_\alpha(z) = \sum_{i_1 \leq \cdots \leq i_n} z_{i_1} \cdots z_{i_n}.$$ 

It is not hard to show that when $|I| \geq n$, the set $\{L_\alpha : \alpha \in \text{Comp}(n)\}$ is a $Q$-basis for $Q_n$ [34, Prop. 7.19.1]. More generally, if $|I| = k$ then the set $\{L_\alpha : \alpha \in \text{Comp}(n), \text{ length}(\alpha) \leq k\}$ is a $Q$-basis for $Q_{\alpha n}$, and $L_\alpha = 0$ if length($\alpha$) > $k$.

If $w = w_1 w_2 \cdots w_n \in S_n$ then let $D(w)$ denote the descent set of $w$, i.e.,

$$D(w) = \{i : w_i > w_{i+1}\}.$$ 

Write co($w$) for the unique composition of $n$ satisfying $S_{\text{co}(w)} = D(w)$. In other words, if co($w$) = ($\alpha_1, \ldots, \alpha_k$), then $w$ has descents exactly at $\alpha_1$, $\alpha_1 + \alpha_2, \ldots, \alpha_1 + \cdots + \alpha_{k-1}$. For any composition $\alpha$ of $n$ it is easy to see that the (possibly infinite) series $L_\alpha(x)$ (where as always $x_i \geq 0$ and $\sum x_i = 1$) is absolutely convergent and therefore represents a nonnegative real number.

The main result on which all our other results depend is the following.

**2.1 Theorem.** Let $w \in S_n$. The probability $\text{Prob}(w)$ that a permutation in $S_n$ chosen from the QS-distribution is equal to $w$ is given by

$$\text{Prob}(w) = L_{\text{co}(w^{-1})}(x).$$

**Proof.** This result is equivalent to [18, Lemma 3.2] (see also [30, Lemma 9.39]). Because of its importance here we present the (easy) proof. The integers $i$ for which $i+1$ appears somewhere to the left of $i$ in $w = w_1 w_2 \cdots w_n$ are just the elements of $D(w^{-1})$. Let $i_1, \ldots, i_n$ be chosen independently from the distribution $x_i$. Let $a_j = i_{w^{-1}(j)}$, i.e., if we write down $i_1 \cdots i_n$ underneath $w_1 \cdots w_n$, then $a_j$ appears below $j$. In order for the sequence $i_1 \cdots i_n$ to standardize to $w$ it is necessary and sufficient that $a_1 \leq a_2 \leq \cdots \leq a_n$ and that $a_j < a_{j+1}$ whenever $j \in D(w^{-1})$. Hence

$$\text{Prob}(w) = \sum_{a_1 \leq a_2 \leq \cdots \leq a_n \atop a_i < a_{i+1} \text{ if } i \in D(w^{-1})} x_{a_1} x_{a_2} \cdots x_{a_n} = L_{\text{co}(w^{-1})}(x). \quad \square$$
For some algebraic results closely related to Theorem 2.1, see [12].

Two special cases of Theorem 2.1 are of particular interest. The first is the $q$-shuffle distribution $U_q$. Let $\text{des}(u)$ denote the number of descents of the permutation $u \in \mathfrak{S}_n$, i.e.,

$$\text{des}(u) = \# \{ i : w(i) > w(i + 1) \}.$$ 

It follows easily from Theorem 2.1 or by a direct argument (see [34, p. 364]) that the probability $\text{Prob}_{U_q}(w)$ that a random permutation in $\mathfrak{S}_n$ chosen from the distribution $U_q$ is equal to $w$ is given by

$$\text{Prob}_{U_q}(w) = \frac{q - \text{des}(w^{-1}) + n - 1}{n} q^{-n}.$$ 

This is just the probability of obtaining $w$ with a $q$-shuffle, as defined in [2, §3], confirming that $U_q$ is just the $q$-shuffle distribution.

For fixed $n$, the uniform distribution $U$ on $\mathfrak{S}_n$ is given by

$$U = \lim_{q \to \infty} U_q.$$ 

Because of this formula many of our results below may be considered as generalizations of the corresponding results for the uniform distribution. These results for the uniform distribution are well-known and classical theorems in the theory of permutation enumeration.

The second interesting special case of Theorem 2.1 is defined by

$$x_i = (1 - t) t^{i-1},$$

where $I = \mathbb{P}$ and $0 \leq t \leq 1$. Given $u = u_1 u_2 \cdots u_n \in \mathfrak{S}_n$, define $\bar{u} = v_1 v_2 \cdots v_n$, where $v_i = n + 1 - u_{n+1-i}$. (Equivalently, $\bar{u} = w_0 u w_0$, where $w_0(i) = n + 1 - i$.) Set $\epsilon(u) = \text{maj}(\bar{u})$, where $\text{maj}$ denotes major index (or greater index) of $u$ [33, p. 216], i.e.,

$$\text{maj}(u) = \sum_{i \in D(u)} i.$$ 

It follows from [34, Lemma 7.19.10] that

$$\text{Prob}(w) = \frac{t^e(w^{-1})(1 - t)^n}{(1 - t)(1 - t^2) \cdots (1 - t^m)}.$$ 

5
The QS-distribution defines a Markov chain (or random walk) $\mathcal{M}_n$ on $\mathfrak{S}_n$ by defining the probability $\text{Prob}(u, uw)$ of the transition from $u$ to $uw$ to be $\text{Prob}(w) = L_{co(w^{-1})}(x)$. The next result determines the eigenvalues of this Markov chain. Equivalently, define

$$\Gamma_n(x) = \sum_{w \in \mathfrak{S}_n} L_{co(w^{-1})}(x)w \in \mathbb{R}\mathfrak{S}_n,$$

where $\mathbb{R}\mathfrak{S}_n$ denotes the group algebra of $\mathfrak{S}_n$ over $\mathbb{R}$. Then the eigenvalues of $\mathcal{M}_n$ are just the eigenvalues of $\Gamma_n(x)$ acting on $\mathbb{R}\mathfrak{S}_n$ by right multiplication.

### 2.2 Theorem

The eigenvalues of $\mathcal{M}_n$ are the power sum symmetric functions $p_\lambda(x)$ for $\lambda \vdash n$. The eigenvalue $p_\lambda(x)$ occurs with multiplicity $n!/z_\lambda$, the number of elements in $\mathfrak{S}_n$ of cycle type $\lambda$.

The above theorem appears implicitly in [16, Thm. 4.4] and more explicitly in [23, Note 5.20]. See also [13, §7] and [34, Exer. 7.96]. It is also a special case of [3, Thm. 1.2], as we now explain.

Bidigare, Hanlon, and Rockmore [3] define a random walk on the set $\mathcal{R}$ of regions of a (finite) hyperplane arrangement $\mathcal{A}$ in $\mathbb{R}^n$ as follows. (Precise definitions of the terms used below related to arrangements may be found in [3]. For further information see [9].) Let $\mathcal{F}$ be the set of faces of $\mathcal{A}$, i.e., the nonempty faces of the closure of the regions of $\mathcal{A}$. Let $w_t$ be a probability distribution on $\mathcal{F}$. Given $R \in \mathcal{R}$ and $F \in \mathcal{F}$, define $F \cdot R$ to be that region $R'$ of $\mathcal{A}$ closest to $R$ that has $F$ as a face. (There is always a unique such region $R'$.) Now given a region $R$, choose $F \in \mathcal{F}$ with probability $w_t(F)$ and move to the region $F \cdot R$.

Consider the special case when $\mathcal{A}$ is the braid arrangement $\mathcal{B}_n$, i.e., the set of hyperplanes $u_i = u_j$ for $1 \leq i < j \leq n$, where $u_1, \ldots, u_n$ are the coordinates in $\mathbb{R}^n$. The set of regions of $\mathcal{B}_n$ can be identified in a natural way with $\mathfrak{S}_n$, viz., identify the region $u_{a_1} < u_{a_2} < \cdots < u_{a_n}$ with the permutation $w$ given by $w(a_i) = i$. The faces of $\mathcal{B}_n$ correspond to ordered set partitions $\pi$ of $1, 2, \ldots, n$, i.e., $\pi = (B_1, \ldots, B_k)$ where $B_i \neq \emptyset$, $B_i \cap B_j = \emptyset$ for $i \neq j$, and $\cup B_i = \{1, 2, \ldots, n\}$; viz., $\pi$ corresponds to the face defined by $u_r = u_s$ if $r, s \in B_i$ for some $i$, and $u_r < u_s$ if $r \in B_i$ and $s \in B_j$ with $i < j$. Define the
type of $\pi$ to be the composition
\[ \text{type}(\pi) = (\#B_1, \ldots, \#B_k) \in \text{Comp}(n). \]

Given $\alpha = (\alpha_1, \ldots, \alpha_k) \in \text{Comp}(n)$, define the monomial quasisymmetric function $M_\alpha(z)$ by
\[ M_\alpha(z) = \sum_{1 \leq i_1 < \cdots < i_k} z_{i_1}^{\alpha_1} \cdots z_{i_k}^{\alpha_k}. \]

(See [34, (7.87)].) Now if $F$ is a face of $\mathcal{B}_n$ corresponding to the ordered partition $\pi$, then set
\[ \text{wt}(F) = M_{\text{type}[\pi]}(x). \]

It is easy to see that $\text{wt}$ is a probability distribution on the set of faces of $\mathcal{B}_n$; in fact, for indeterminates $(z_i)_{i \in I}$ we have
\[ \sum_\pi M_{\text{type}[\pi]}(z) = \left( \sum z_i \right)^n. \]

Thus $\text{wt}$ defines a Bidigare-Hanlon-Rockmore random walk on the regions of $\mathcal{B}_n$. Let $P(R, R')$ denote the probability of moving from a region $R$ to a region $R'$.

**2.3 Theorem.** Let $R$ correspond to $u \in \mathfrak{S}_n$ and $R'$ to $uv$. Then
\[ P(R, R') = I_{\text{co}(w^{-1})}(x). \]
Hence the Bidigare-Hanlon-Rockmore random walk just defined on the regions of $\mathcal{B}_n$ is isomorphic to the random walk on $\mathfrak{S}_n$ defined by the QS-distribution.

**Proof.** By symmetry we may suppose that $u = 12 \cdots n$, the identity permutation. The faces $F$ of $R'$ for which $F \cdot R = R'$ correspond to those ordered partitions $\pi$ obtained from $w^{-1} = w_1 \cdots w_n$ by drawing bars between certain pairs $w_i$ and $w_{i+1}$, where there must be a bar when $i \in D(w^{-1})$. The elements contained between the consecutive bars (including a bar before $w_1$ and after $w_n$) form the blocks of $\pi$, read from left-to-right. For instance, if $w^{-1} = 582679134$, then one way of drawing bars is $[58][267][9][1][34]$, yielding the ordered partition $\pi = (\{5, 8\}, \{2, 6, 7\}, \{9\}, \{1\}, \{3, 4\})$. It follows that
\[ P(R, R') = \sum_\beta M_\beta(x), \]
where \( \beta \) runs over all compositions refining \( \text{co}(w^{-1}) \). From the definition of \( L_{\alpha}(x) \) we have
\[
\sum_{\beta} M_{\beta}(x) = L_{\text{co}(w^{-1})}(x),
\]
and the proof follows. \( \square \)

The next result determines the convolution of the QS\((x)\)-distribution with the QS\((y)\)-distribution. We write \( xy \) for the variables \( x_i y_j \) in the order \( x_i y_j < x_r y_s \) if either \( i < r \) or \( i = r \) and \( j < s \).

### 2.4 Theorem
Suppose that a permutation \( u \in \mathfrak{S}_n \) is chosen from the QS\((x)\)-distribution, and a permutation \( v \in \mathfrak{S}_n \) is chosen from the QS\((y)\)-distribution. Let \( w \in \mathfrak{S}_n \). Then the probability that \( uv = w \) is equal to \( L_{\text{co}(w^{-1})}(xy) \). In other words, if \( * \) denotes convolution then
\[
\text{QS}(x) * \text{QS}(y) = \text{QS}(xy).
\]
Equivalently, in the ring \( R\mathfrak{S}_n \) we have
\[
\Gamma_n(x)\Gamma_n(y) = \Gamma_n(xy).
\]

Theorem 2.4 is equivalent to [23, (62)] [30, Thm. 9.37]. For the \( U_q \)-distribution it was also proved in [2, Lemma 1], and this proof can be easily extended to prove Theorem 2.4 in its full generality.

### 3 Enumerative properties of the QS-distribution

Suppose that \( X \) is any subset of \( \mathfrak{S}_n \). Let \( w \in \mathfrak{S}_n \) be chosen from the QS-distribution. It follows from Theorem 2.1 that the probability that \( w \in X \) is given by
\[
\text{Prob}(w \in X) = \sum_{u \in X} L_{\text{co}(u^{-1})}(x). \tag{3}
\]
There are many known instances of the set \( X \subset \mathfrak{S}_n \) for which the right-hand side of (3) can be explicitly evaluated. Most of this section will be devoted to some examples of this type.
Our first result involves the symmetric function $l_n(z)$ defined by

$$l_n(z) = \frac{1}{n} \sum_{d|n} \mu(d) p_{d^j}^n(z),$$

where $\mu$ denotes the usual number-theoretic Möbius function. More information on $l_n(z)$ may be found in [30, Ch. 8] or [34, Exer. 7.88–7.89].

3.1 Theorem. Let $w$ be a random permutation in $\mathfrak{S}_n$, chosen from the QS-distribution. The probability $\text{Prob}(\rho(w) = \lambda)$ that $w$ has cycle type $\lambda = \langle 1^{m_1} 2^{m_2} \cdots \rangle \vdash n$ (i.e., $m_i$ cycles of length $i$) is given by

$$\text{Prob}(\rho(w) = \lambda) = \prod_{i \geq 1} h_m[i_i](x),$$

where brackets denote plethysm [26, §I.8][34, p. 447].

The proof of Theorem 3.1 follows from [30, Thm. 8.23 and Thm. 9.41(a)]. In the special case of the $U_q$-distribution, the result appears in [11, Thm. A].

As above let $\text{maj}(w)$ denote the major index of $w \in \mathfrak{S}_n$, and let $\text{inv}(w)$ denote the number of inversions of $w$ [33, pp. 20–21], i.e.,

$$\text{inv}(w) = \# \{(i, j) : 1 \leq i < j \leq n, \ w(i) > w(j)\}.$$ 

Let $I_n(j)$ (respectively, $M_n(j)$) denote the probability that $w \in \mathfrak{S}_n$ satisfies $\text{inv}(w) = j$ (respectively, $\text{maj}(w) = j$) under the QS-distribution.

3.2 Theorem. We have

$$\sum_{n \geq 0} \sum_{j \geq 0} \frac{M_n(j) t^j z^n}{(1-t)(1-t^2) \cdots (1-t^n)} = \prod_{i \geq 1} \prod_{j \in \mathbb{I}} \left(1 - t^{i-1} x_j z \right)^{-1}. \quad (4)$$

Proof. The standardization $w$ of a sequence $i = i_1 i_2 \cdots i_n$ satisfies $\text{maj}(w) = \text{maj}(\hat{i})$ and $\text{inv}(w) = \text{inv}(\hat{i})$. Hence by definition of the QS-distribution we have

$$\sum_j M_n(j) t^j = \sum_{\hat{i}=i_1 \cdots i_n} t^{\text{maj}(\hat{i})} x_{i_1} \cdots x_{i_n},$$

$$\sum_j I_n(j) t^j = \sum_{\hat{i}=i_1 \cdots i_n} t^{\text{inv}(\hat{i})} x_{i_1} \cdots x_{i_n}.$$
Thus if
\[ F_\lambda(t) = \sum_{v} t^{\text{maj}(v)} \]
\[ G_\lambda(t) = \sum_{v} t^{\text{inv}(v)}, \]
where \( v \) ranges over all permutations of the multiset \( \{1^{\lambda_1}, 2^{\lambda_2}, \ldots\} \), then
\[ \sum_j M_n(j)t^j = \sum_{\lambda \vdash n} F_\lambda(t)m_\lambda(x) \]
\[ \sum_j I_n(j)t^j = \sum_{\lambda \vdash n} G_\lambda(t)m_\lambda(x), \]
where \( m_\lambda(x) \) denotes a monomial symmetric function. But it is known [32, (45), p. 97][33, Prop. 1.3.17] that
\[ F_\lambda(t) = G_\lambda(t) = \left[ \begin{array}{c} n \\ \lambda_1, \lambda_2, \ldots \end{array} \right]_t, \]
a \( q \)-multinomial coefficient in the variable \( t \). Hence \( M_n(j) = I_n(j) \).

Now let \( h_\lambda(x) \) denote a complete homogeneous symmetric function [34, §7.5]. It is easy to see [34, Prop. 7.8.3] that
\[ h_\lambda(1, t, t^2, \ldots) = \prod_i \left( \frac{1}{1-t} \right) \left( \frac{1}{1-t^2} \right) \cdots \left( \frac{1}{1-t^n} \right) \cdot \]
Hence
\[ \sum_j M_n(j)t^j = (1-t) \cdots (1-t^n) \sum_{\lambda \vdash n} h_\lambda(1, t, t^2, \ldots)m_\lambda(x). \]
Equation (4) then follows immediately from the expansion [34, (7.10)] of the Cauchy product \( \prod_{i,j}(1-x_iy_j)^{-1} \). Equation (4) is also equivalent to [17, (78)]. \( \square \)

Other expansions of the Cauchy product lead to further formulas for \( \sum_j M_n(j)t^j \). In particular, from [34, (7.20), Thm. 7.12.1, and Cor. 7.21.3]
there follows (using notation from [34])
\[
\sum_j M_n(j) t^j = \sum_{\lambda\vdash n} t^{b(\lambda)} \frac{(1-t) \cdots (1-t^n)}{\prod_{u \in \lambda} (1-t^{h(u)})} s_\lambda(x)
\]
\[
= \sum_{\lambda\vdash n} \frac{(1-t) \cdots (1-t^n)}{(1-t^{\lambda_1}) \cdots (1-t^{\lambda_k})} z_\lambda \lambda_1^{-1} p_\lambda(x). \tag{5}
\]

A famous result of MacMahon asserts that
\[
\# \{w \in \mathfrak{S}_n : \text{maj}(w) = j\} = \# \{w \in \mathfrak{S}_n : \text{inv}(w) = j\}.
\]

Since the uniform distribution $U$ on $\mathfrak{S}_n$ satisfies $U = \lim_{q \to \infty} U_q$, it follows that Theorem 3.2 is a generalization of MacMahon’s result. In fact, the proof we have given of Theorem 3.2 shows that it is equivalent to the equidistribution of maj and inv on the set of permutations of any finite multiset (where the underlying set is linear ordered). This result was also known, at least implicitly, to MacMahon. For further references and some generalizations, see [5].

In principle one can use Theorem 3.2 to derive the moments of $\text{maj}(w)$ (or equivalently $\text{inv}(w)$). Let $E$ denote expectation with respect to the QS-distribution and $E_U$ expectation with respect to the uniform distribution. The next result obtains the first two moments of $\text{maj}(w)$, stated for simplicity as the expectations of $\text{maj}(w)$ and $\text{maj}(w)(\text{maj}(w) - 1)$ (instead of $\text{maj}(w)^2$).

**3.3 Corollary.** Choose $w \in \mathfrak{S}_n$ from the QS-distribution. Then
\[
E(\text{maj}(w)) = E_U(\text{maj}(w)) - \frac{1}{2} \binom{n}{2} p_2(x) \tag{6}
\]
\[
E(\text{maj}(w)(\text{maj}(w) - 1)) = E_U(\text{maj}(w)(\text{maj}(w) - 1)) \tag{7}
\]
\[
-3 \binom{n+1}{4} p_2(x) + \frac{4}{3} \binom{n}{3} p_3(x) + \frac{3}{2} \binom{n}{4} p_2(x)^2. \tag{8}
\]

**Note.** It is easy to see (e.g. [22, p. 16]) that
\[
E_U(\text{maj}(w)) = \frac{1}{2} \binom{n}{2},
\]
\[
E_U(\text{maj}(w)(\text{maj}(w) - 1)) = \frac{n(n-1)(n-2)(9n+13)}{144}.
\]
Proof of Corollary 3.3. Let

$$\Lambda_n(t) = \sum_j M_n(j)t^j.$$ 

Then

$$E(\text{maj}(w)) = \Lambda'_n(1).$$

When we differentiate (5) term-by-term and set $t = 1$, the only surviving terms will be from $\lambda = \langle 1^n \rangle$ and $\lambda = \langle 2^{n-2} \rangle$. Hence

$$E(\text{maj}(w)) = \frac{d}{dt} \left[ \frac{(1-t) \cdots (1-t^n)}{(1-t)^n} \frac{1}{n!} p_1(x)^n \right]_{t=1} + \frac{(1-t) \cdots (1-t^n)}{(1-t^2)(1-t)^{n-2}} \frac{1}{2 \cdot (n-2)!} p_1(x)^{n-2} p_2(x) \right]_{t=1}. \quad (9)$$

Now $p_1(x) = \sum x_i = 1$ and

$$\frac{d}{dt} \left[ \frac{(1-t) \cdots (1-t^n)}{(1-t)^n} \frac{1}{n!} \right]_{t=1} = E_U(\text{maj}(w)).$$

It is routine to compute the second term on the right-hand side of (9), obtaining equation (6).

A similar computation using $E(\text{maj}(w)(\text{maj}(w) - 1)) = \Lambda''_n(1)$ yields (8); we omit the details. □

It is clear from the above proof that for general $k \geq 1$, $E(\text{maj}(w)^k)$ is a linear combination, whose coefficients are polynomials in $n$, of power sum symmetric functions $p_\lambda(x)$ where $\lambda \vdash n$ and $\ell(\lambda) \geq n - k$. Here $\ell(\lambda)$ denotes the length (number of parts) of $\lambda$.

We next consider the relationship between the QS-distribution and the Robinson-Schensted-Knuth (RSK) algorithm [34, §7.11]. Recall that this algorithm associates a pair $(T, T')$ of standard Young tableaux (SYT) of the same shape $\lambda \vdash n$ with a permutation $w \in \mathfrak{S}_n$. We call $T'$ the recording tableau of $w$, denoted $\text{rec}(w)$. The shape $\lambda$ of $T'$ is also called the shape of $w$, denoted $\text{sh}(w) = \lambda$. 

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3.4 Theorem. Choose $w \in \mathfrak{S}_n$ from the QS-distribution. Let $T$ be an SYT of shape $\lambda \vdash n$. Then the probability that $\text{rec}(w) = T$ is given by

$$\text{Prob}(\text{rec}(w) = T) = s_\lambda(x),$$

where $s_\lambda$ denotes a Schur function.

Proof. Let

$$X_T = \{w \in \mathfrak{S}_n : \text{rec}(w) = T\}.$$

It follows from [34, Thm. 7.19.2 and Lemma 7.23.1] that

$$\sum_{w \in X_T} L_{\text{co}(w^{-1})}(x) = s_\lambda(x).$$

The proof now follows from equation (3). □

3.5 Corollary. Choose $w \in \mathfrak{S}_n$ from the QS-distribution, and let $\lambda \vdash n$. Then

$$\text{Prob}(\text{sh}(w) = \lambda) = f^\lambda s_\lambda(x),$$

where $f^\lambda$ denotes the number of SYT of shape $\lambda$ (given explicitly by the Frame-Robinson-Thrall hook-length formula [34, Cor. 7.21.6]).

Proof. There are $f^\lambda$ SYT’s $T$ of shape $\lambda$, and by Theorem 3.4 they all have probability $s_\lambda(x)$ of being the recording tableau of $w$. □

Note. If the RSK algorithm associates $(T, T')$ with $w \in \mathfrak{S}_n$, then call $T$ the insertion tableau of $w$, denoted $\text{ins}(w)$. Since $\text{ins}(w) = \text{rec}(w^{-1})$ [34, Thm. 7.13.1], it follows that

$$\text{Prob}(\text{ins}(w) = T) = f^\lambda L_{\text{co}(w)}(x).$$

Note. The probability distribution $\text{Prob}(\lambda) = f^\lambda s_\lambda(x)$ on the set of all partitions $\lambda$ of all nonnegative integers is a specialization of the $z$-measure of Borodin and Olshanski, as surveyed in [7] (see also [29]).

It is easy to give an expression for the probability that a permutation $w \in \mathfrak{S}_n$ chosen from the QS-distribution has a fixed descent set $S \subseteq \{1, 2, \ldots, n-
In general, if \( \lambda \) and \( \mu \) are partitions with \( \mu \subset \lambda \) (i.e., \( \mu_i \leq \lambda_i \) for all \( i \)), then \( \lambda/\mu \) denotes the skew shape obtained by removing the diagram of \( \mu \) from that of \( \lambda \) [34, p. 309]. One can then define (see [34, Def. 7.10.1]) the skew Schur function \( s_{\lambda/\mu}(x) \). A border strip (or rim hook or ribbon) is a connected skew shape with no \( 2 \times 2 \) square [34, p. 245]. Let \( \alpha = (\alpha_1, \ldots, \alpha_k) \in \text{Comp}(n) \). Then there is a unique border strip \( B_\alpha \) (up to translation) with \( \alpha_i \) squares in row \( i \). For instance, the border strip \( B_{(3,2,1,2,1)} \) looks like

\[
\begin{array}{cccccc}
\square & & & & \\
& \square & & & & \\
& & \square & & & \\
& & & \square & & \\
& & & & \square & \\
\end{array}
\]

Some special properties of border strip Schur functions \( s_{B_\alpha} \) are discussed in [34, §7.23].

**3.6 Theorem.** Let \( w \) be a random permutation in \( \mathcal{S}_n \), chosen from the QS-distribution. Let \( \alpha = (\alpha_1, \ldots, \alpha_k) \in \text{Comp}(n) \), and recall that \( S_\alpha = \{\alpha_1, \alpha_1 + \alpha_2, \ldots, \alpha_1 + \cdots + \alpha_{k-1}\} \subseteq \{1, 2, \ldots, n-1\} \). Then the probability that \( w \) has descent set \( S_\alpha \) is given by

\[
\text{Prob}(D(w) = S_\alpha) = s_{B_\alpha}(x).
\]

**Proof.** Immediate from Theorem 2.1 and the fact [34, Cor. 7.23.4] that

\[
s_{B_\alpha} = \sum_{\substack{w \in \mathcal{S}_n \\
\alpha = \text{co}(w) \}} L_{\text{co}(w-1)}. \quad \square
\]

4 Longest increasing subsequences.

It is a fundamental result of Schensted [31] (see also [34, Cor. 7.23.11]) that if \( \text{sh}(w) = \lambda = (\lambda_1, \lambda_2, \ldots) \), then \( \lambda_1 \) is the length \( \text{is}(w) \) of the longest increasing subsequence of \( w \). (There is a generalization due to Greene [19][34, Thm. 7.23.13 or A1.1.1] that gives a similar interpretation of any \( \lambda_i \).) Thus
the expected value $E_U(n)$ of $is(w)$ for $w \in \mathcal{S}_n$ under the uniform distribution is given by

$$E_U(n) = E_U(is(w)) = \frac{1}{n!} \sum_{\lambda \vdash n} \lambda_1 (f^\lambda)^2.$$  

(See [34, Exer. 7.109(b)].) It was shown by Vershik and Kerov [35] that $E_U(n) \sim 2\sqrt{n}$. Later Baik et al. [1] obtained a much stronger result, viz., the exact distribution (suitably normalized) of $\lambda_1$ (or equivalently $is(w)$) in the limit $n \to \infty$. (This result was extended to any $\lambda_i$ in [8][28][21].) We can ask whether similar results hold for the QS-distribution. Many results along these lines appear in [4][6, Ex. 2][8][27][29]. Here we make some elementary comments pointing in a different direction from these papers.

It is clear from Corollary 3.5 that

$$E(is(w)) = \sum_{\lambda \vdash n} \lambda_1 f^\lambda s_\lambda(x), \quad (10)$$

where as usual $E$ denotes expectation with respect to the QS-distribution. This formula can be made more explicit in the special case of the distribution $U_q$. Identify $\lambda \vdash n$ with its Young diagram, and let $c(u)$ denote the content of the square $u \in \lambda$ as defined in [34, p. 373].

4.1 Theorem. We have

$$E_{U_q}(is(w)) = \frac{1}{n!} \sum_{\lambda \vdash n} \lambda_1 (f^\lambda)^2 \prod_{u \in \lambda} (1 + q^{-1}c(u)). \quad (11)$$

Proof. Let $s_\lambda(1^q)$ denote $s_\lambda(x)$ where $x_1 = \cdots = x_q = 1$ and $x_i = 0$ for $i > q$. It is well-known (e.g. [34, Cor. 7.21.4 and 7.21.6]) that

$$s_\lambda(1^q) = \frac{f^\lambda}{n!} \prod_{u \in \lambda} (q + c(u)).$$

Since $s_\lambda(x)$ is homogeneous of degree $n$, the proof follows from equation (10). \□

Note that $E_{U_q}(is(w))$ is a polynomial in $q^{-1}$, which can be written

$$E_{U_q}(is(w)) = E_U(is(w)) + \frac{1}{n!} \sum_{\lambda \vdash n} \lambda_1 (f^\lambda)^2 \left( \sum_{u \in \lambda} c(u) \right) \frac{1}{q} + \cdots. \quad (12)$$
Let us mention that

\[ \sum_{u \in \lambda} c(u) = \sum i(\lambda'_i - \lambda_i) = \sum \left( \frac{\lambda_i}{2} \right) - \sum \left( \frac{\lambda'_i}{2} \right), \]

where \( \lambda' = (\lambda'_1, \lambda'_2, \ldots) \) denotes the conjugate partition to \( \lambda \). If we regard \( n \) as fixed, then equation (12) shows the rate at which \( E_U(q \text{is}(w)) \) converges to \( E_U(\text{is}(w)) \) as \( q \to \infty \). It is therefore of interest to investigate the asymptotic behavior of the coefficient

\[ F_1(n) = \frac{1}{n!} \sum_{\lambda \vdash n} \lambda_1 \left( f^\lambda \right)^2 \left( \sum_{u \in \lambda} c(u) \right) \]

as \( n \to \infty \) (as well as the coefficient of \( q^{-i} \) for \( i > 1 \)). Numerical evidence suggests that \( F_1(n)/n \) is a slowly increasing function of \( n \). The largest value of \( n \) for which we have done the computation is \( n = 47 \), giving \( F_1(47)/47 \approx 0.6991 \). Eric Rains suggests that there is some reason to suspect that \( F(n)/n \) grows like \( n^{1/6} \).

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